Imprecise probability models and prior-data conflict

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11.12.2007

prior-data conflict

LUCK-models

$$p(\vartheta) \propto \exp\left\{n^{(0)} \left[\langle \psi, y^{(0)} \rangle - \mathbf{b}(\psi) \right]\right\}$$

and

$$p(\vartheta \mid w) \propto \exp \left\{ n^{(1)} \left[\langle \psi, y^{(1)} \rangle - \mathbf{b}(\psi) \right] \right\},$$

with

$$n^{(1)} = n^{(0)} + q$$
 and $y^{(1)} = \frac{n^{(0)}y^{(0)} + \tau(w)}{n^{(0)} + q}$.

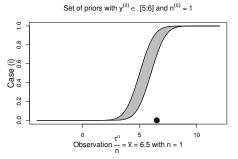
imprecise priors

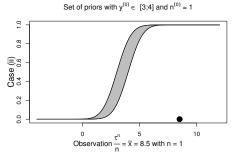
Just like in the IDM: set $\mathcal{Y}^{(0)} \subset \mathcal{Y}$ and $n^{(0)}$ fix.

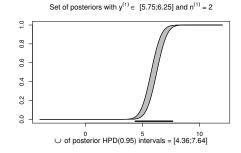
$$\mathcal{Y}^{(1)} = \left\{ \left. \frac{n^{(0)}y^{(0)} + \tau(w)}{n^{(0)} + n} \, \right| \, y^{(0)} \in \mathcal{Y}^{(0)}
ight\} \subset \mathcal{Y} \, .$$

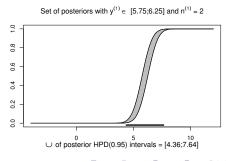
 $\mathcal{Y}^{(1)}$ can actually be seen as a shifted and rescaled version of $\mathcal{Y}^{(0)}$:

$$\mathcal{Y}^{(1)} = \frac{n^{(0)}}{n^{(0)} + n} \cdot \mathcal{Y}^{(0)} + \frac{n}{n^{(0)} + n} \cdot \frac{1}{n} \sum_{i=1}^{n} \tau(w_i),$$







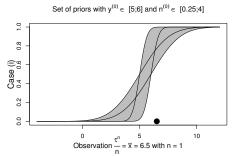


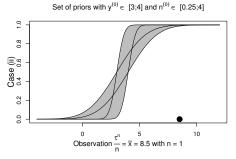
Why that?

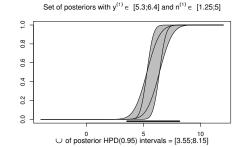
$$\overline{y}^{(1)} - \underline{y}^{(1)} = \frac{n^{(0)} \left(\overline{y}^{(0)} - \underline{y}^{(0)} \right)}{n^{(0)} + n} \text{ does not depend on } \tau^n(x)!$$

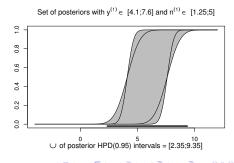
Idea: vary $n^{(0)}$ in some set $\mathcal{N}^{(0)}$ as well!











Why does it work?

$$\underline{y}^{(1)} = \begin{cases} \frac{\overline{n}^{(0)}\underline{y}^{(0)} + \tau^{n}(x)}{\overline{n}^{(0)} + n} & \text{if } \frac{\tau^{n}(x)}{n} \ge \underline{y}^{(0)} \\ \frac{\underline{n}^{(0)}\underline{y}^{(0)} + \tau^{n}(x)}{\underline{n}^{(0)} + n} & \text{if } \frac{\tau^{n}(x)}{n} < \underline{y}^{(0)} & \iff \text{prior-data conflict} \end{cases}$$

$$\overline{y}^{(1)} = \begin{cases} \frac{\overline{n}^{(0)}\overline{y}^{(0)} + \tau^{n}(x)}{\overline{n}^{(0)} + n} & \text{if } \frac{\tau^{n}(x)}{n} \leq \overline{y}^{(0)} \\ \frac{\underline{n}^{(0)}\overline{y}^{(0)} + \tau^{n}(x)}{\underline{n}^{(0)} + n} & \text{if } \frac{\tau^{n}(x)}{n} > \overline{y}^{(0)} & \iff \text{prior-data conflict} \end{cases}$$

Now:

$$\overline{y}^{(1)} - \underline{y}^{(1)} = \frac{\overline{n}^{(0)}(\overline{y}^{(0)} - \underline{y}^{(0)})}{\overline{n}^{(0)} + n} + \Delta \left(\frac{\tau^{n}(x)}{n}; \underline{y}^{(0)}, \overline{y}^{(0)} \right) \frac{n(\overline{n}^{(0)} - \underline{n}^{(0)})}{(\overline{n}^{(0)} + n)(\underline{n}^{(0)} + n)},$$

where

$$\Delta\left(\frac{\tau^n(x)}{n}; \underline{y}^{(0)}, \overline{y}^{(0)}\right) = \inf\left\{\left|\frac{\tau^n(x)}{n} - y^{(0)}\right| : \underline{y}^{(0)} \le y^{(0)} \le \overline{y}^{(0)}\right\}$$

is the distance between the prior interval $[\underline{y}^{(0)}; \overline{y}^{(0)}]$ and the observation $\frac{\tau^n(x)}{n}$.



Why then not choose $\sup \mathcal{N}^{(0)} = n^{(0)}$?

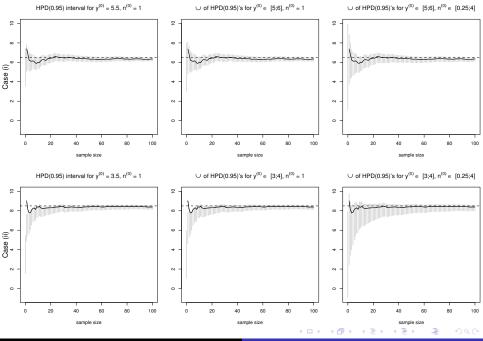
The factor to $\Delta(\)$ gets maximal if $n=\sqrt{\underline{n}^{(0)}\overline{n}^{(0)}}$.

$$\iff \overline{y}^{(1)} - y^{(1)}$$
 maximal for fixed $\Delta(\)$.

 \iff Same weight on the prior and on the sample.

 $\iff \sqrt{\underline{n}^{(0)}\overline{n}^{(0)}}$ is the 'global' prior strength to be compared to $n^{(0)}$.

For $n \longrightarrow \infty$, it doesn't matter, as $\frac{\overline{n}^{(1)} - \underline{n}^{(1)}}{\overline{n}^{(1)}} \longrightarrow 0$:



\underline{P} and \overline{P} for classical confidence intervals

$$\mathsf{CI} = \bar{x} \pm z_{\frac{1+\gamma}{2}} \frac{1}{n}$$
 (we have $\sigma^2 = 1$)

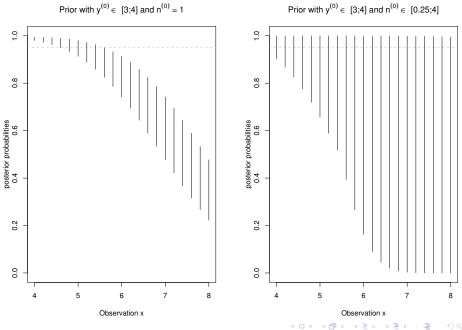
$$[\underline{P}(\{\mu \in \mathsf{CI}\}); \overline{P}(\{\mu \in \mathsf{CI}\})]$$

should tend to vacuous the more severe the prior-data conflict is.

Probability weight of area around \bar{x} should be very insecure:

- if data 'is right', CI should have high probability weight
- if prior 'is right', CI should have low probability weight





dataset: Mietspiegel

Summary