

# Nonparametric Bayesian System Reliability using Sets of Priors

Gero Walter<sup>1</sup>, Louis Aslett<sup>2</sup>, Frank Coolen<sup>3</sup>

<sup>1</sup>Eindhoven University of Technology, Eindhoven, NL

<sup>2</sup>University of Oxford, Oxford, UK

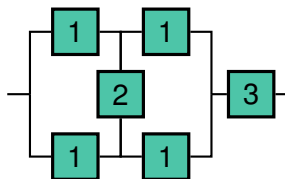
<sup>3</sup>Durham University, Durham, UK

[g.m.walter@tue.nl](mailto:g.m.walter@tue.nl)



Hannover 2016-12-08

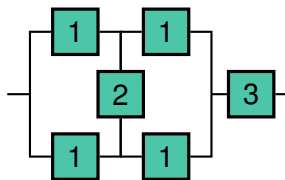
We want to learn about the system reliability  
 $R_{\text{sys}}(t) = P(T_{\text{sys}} > t)$  (system survival function)  
based on

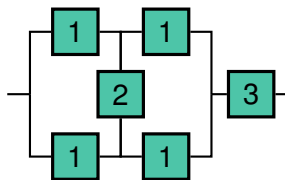


We want to learn about the system reliability  
 $R_{\text{sys}}(t) = P(T_{\text{sys}} > t)$  (system survival function)  
based on

▶ component test data:

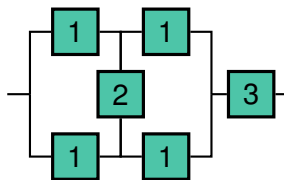
$n_k$  failure times for components of type  $k$ ,  
 $k = 1, \dots, K$





We want to learn about the system reliability  
 $R_{\text{sys}}(t) = P(T_{\text{sys}} > t)$  (system survival function)  
based on

- ▶ component test data:  
 $n_k$  failure times for components of type  $k$ ,  
 $k = 1, \dots, K$
- ▶ cautious assumptions  
on component reliability:  
expert information,  
e.g. from maintenance managers and staff



We want to learn about the system reliability  
 $R_{\text{sys}}(t) = P(T_{\text{sys}} > t)$  (system survival function)  
based on

- ▶ component test data:  
 $n_k$  failure times for components of type  $k$ ,  
 $k = 1, \dots, K$
- ▶ cautious assumptions  
on component reliability:  
expert information,  
e.g. from maintenance managers and staff

How to combine these two information sources?

expert info + data → complete picture

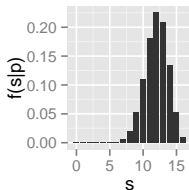
expert info	+	data	→	complete picture
prior distribution	+	sample distribution	→	posterior distribution
$f(p)$	×	$f(s   p)$	$\propto$	$f(p   s)$
				▶ Bayes' Rule

expert info	+	data	→ complete picture
prior distribution	+	sample distribution	→ posterior distribution
$f(p)$	×	$f(s   p)$	$\propto f(p   s)$

► Bayes' Rule

## Binomial distribution

$$s | p \sim \text{Binomial}(n, p)$$





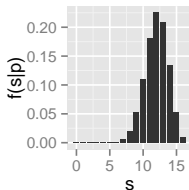
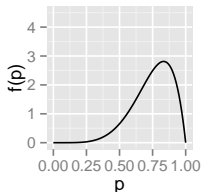
expert info	+	data	→ complete picture
prior distribution	+	sample distribution	→ posterior distribution
$f(p)$	×	$f(s   p)$	$\propto f(p   s)$
			▶ Bayes' Rule

Beta prior

Binomial  
distribution

$$p \sim \text{Beta}(\alpha^{(0)}, \beta^{(0)})$$

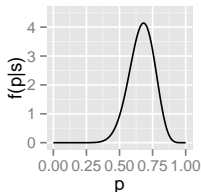
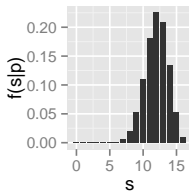
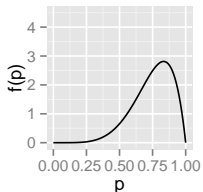
$$s | p \sim \text{Binomial}(n, p)$$



expert info	+	data	→ complete picture
prior distribution	+	sample distribution	→ posterior distribution
$f(p)$	×	$f(s   p)$	$\propto f(p   s)$
Beta prior		Binomial distribution	Beta posterior
$p \sim \text{Beta}(\alpha^{(0)}, \beta^{(0)})$		$s   p \sim \text{Binomial}(n, p)$	$p   s \sim \text{Beta}(\alpha^{(n)}, \beta^{(n)})$

► Bayes' Rule

► conjugacy



expert info + data  $\rightarrow$  complete picture

prior distribution + sample distribution  $\rightarrow$  posterior distribution

$$f(p) \times f(s | p) \propto f(p | s)$$

► Bayes' Rule

Beta prior

Binomial  
distribution

Beta posterior

► conjugacy

$$p \sim \text{Beta}(\alpha^{(0)}, \beta^{(0)})$$

$$s | p \sim \text{Binomial}(n, p)$$

$$p | s \sim \text{Beta}(\alpha^{(n)}, \beta^{(n)})$$

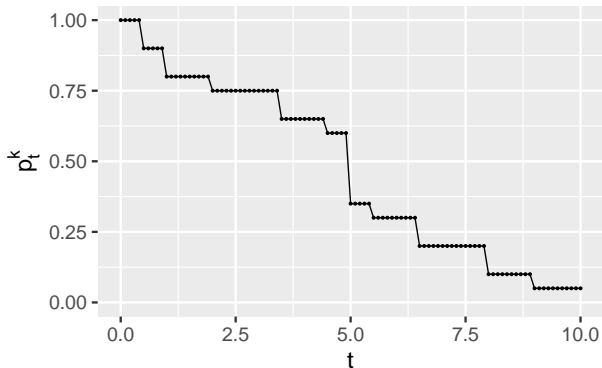
- conjugate prior makes learning about parameter tractable, just update hyperparameters:  $\alpha^{(0)} \rightarrow \alpha^{(n)}, \beta^{(0)} \rightarrow \beta^{(n)}$
- closed form for many inferences, e.g.  $E[p | s] = \frac{\alpha^{(n)}}{\alpha^{(n)} + \beta^{(n)}}$

Functioning probability  $p_t^k$  of **k** for each time  $t \in \mathcal{T} = \{t'_1, t'_2, \dots\}$

- ▶ discrete component reliability function  $R^k(t) = p_t^k, t \in \mathcal{T}$ .

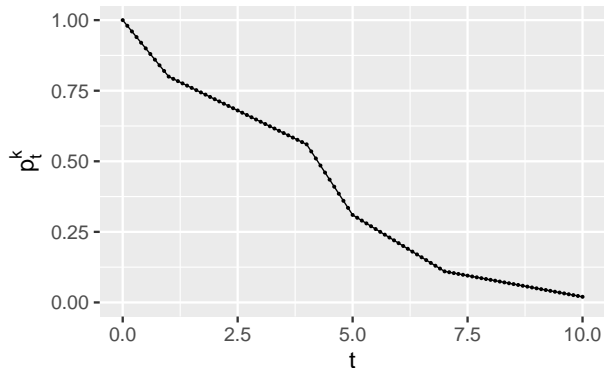
Functioning probability  $p_t^k$  of **k** for each time  $t \in \mathcal{T} = \{t'_1, t'_2, \dots\}$

▶ discrete component reliability function  $R^k(t) = p_t^k, t \in \mathcal{T}$ .



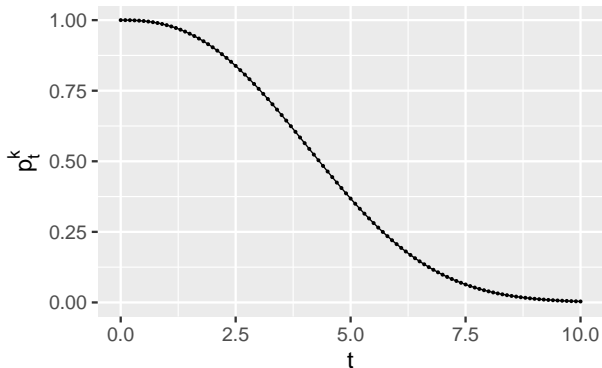
Functioning probability  $p_t^k$  of **k** for each time  $t \in \mathcal{T} = \{t'_1, t'_2, \dots\}$

▶ discrete component reliability function  $R^k(t) = p_t^k, t \in \mathcal{T}$ .



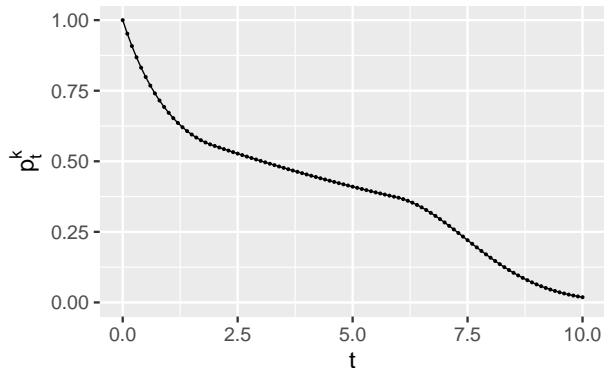
Functioning probability  $p_t^k$  of **k** for each time  $t \in \mathcal{T} = \{t'_1, t'_2, \dots\}$

▶ discrete component reliability function  $R^k(t) = p_t^k, t \in \mathcal{T}$ .



Functioning probability  $p_t^k$  of **k** for each time  $t \in \mathcal{T} = \{t'_1, t'_2, \dots\}$

- ▶ discrete component reliability function  $R^k(t) = p_t^k, t \in \mathcal{T}$ .





Functioning probability  $p_t^k$  of **k** for each time  $t \in \mathcal{T} = \{t'_1, t'_2, \dots\}$

- ▶ discrete component reliability function  $R^k(t) = p_t^k, t \in \mathcal{T}$ .

use Bayesian inference to estimate  $p_t^k$ 's:

Functioning probability  $p_t^k$  of **k** for each time  $t \in \mathcal{T} = \{t'_1, t'_2, \dots\}$

▶ discrete component reliability function  $R^k(t) = p_t^k, t \in \mathcal{T}$ .

use Bayesian inference to estimate  $p_t^k$ 's:

▶ failure times  $\mathbf{t}^k = (t_1^k, \dots, t_{n_k}^k)$  from component test data

number of type  $k$  components functioning at  $t$ :

$$S_t^k | p_t^k \sim \text{Binomial}(p_t^k, n_k)$$

Functioning probability  $p_t^k$  of **k** for each time  $t \in \mathcal{T} = \{t'_1, t'_2, \dots\}$

- ▶ discrete component reliability function  $R^k(t) = p_t^k, t \in \mathcal{T}$ .

use Bayesian inference to estimate  $p_t^k$ 's:

- ▶ failure times  $\mathbf{t}^k = (t_1^k, \dots, t_{n_k}^k)$  from component test data

number of type  $k$  components functioning at  $t$ :

$$S_t^k \mid p_t^k \sim \text{Binomial}(p_t^k, n_k)$$

- ▶ expert knowledge

Beta prior for each  $k$  and  $t$ :

$$p_t^k \sim \text{Beta}(\alpha_{k,t}^{(0)}, \beta_{k,t}^{(0)})$$

Functioning probability  $p_t^k$  of **k** for each time  $t \in \mathcal{T} = \{t'_1, t'_2, \dots\}$

- ▶ discrete component reliability function  $R^k(t) = p_t^k, t \in \mathcal{T}$ .

use Bayesian inference to estimate  $p_t^k$ 's:

- ▶ failure times  $t^k = (t_1^k, \dots, t_{n_k}^k)$  from component test data

number of type  $k$  components functioning at  $t$ :

$$S_t^k | p_t^k \sim \text{Binomial}(p_t^k, n_k)$$

- ▶ expert knowledge

Beta prior for each  $k$  and  $t$ :

$$p_t^k \sim \text{Beta}(\alpha_{k,t}^{(0)}, \beta_{k,t}^{(0)})$$

- ▶ complete picture

Beta posterior for each  $k$  and  $t$ :

$$p_t^k | s_t^k \sim \text{Beta}(\alpha_{k,t}^{(n)}, \beta_{k,t}^{(n)})$$

What if expert information and data tell different stories?

What if expert information and data tell different stories?

## Prior-Data Conflict

- ▶ *informative prior beliefs* and *trusted data* (sampling model correct, no outliers, etc.) are in conflict
- ▶ “[. . .] the prior [places] its mass primarily on distributions in the sampling model for which the observed data is surprising” (Evans and Moshonov 2006)
- ▶ there are not enough data to overrule the prior

What if expert information and data tell different stories?

- ▶ reparametrisation helps to understand effect of prior-data conflict:

What if expert information and data tell different stories?

- ▶ reparametrisation helps to understand effect of prior-data conflict:

$$n^{(0)} = \alpha^{(0)} + \beta^{(0)}, \quad y^{(0)} = \frac{\alpha^{(0)}}{\alpha^{(0)} + \beta^{(0)}}, \quad \text{which are updated as}$$

$$n^{(n)} = n^{(0)} + n, \quad y^{(n)} = \frac{n^{(0)}}{n^{(0)} + n} y^{(0)} + \frac{n}{n^{(0)} + n} \cdot \frac{s}{n}$$



What if expert information and data tell different stories?

- ▶ reparametrisation helps to understand effect of prior-data conflict:

$$n^{(0)} = \alpha^{(0)} + \beta^{(0)}, \quad y^{(0)} = \frac{\alpha^{(0)}}{\alpha^{(0)} + \beta^{(0)}}, \quad \text{which are updated as}$$
$$n^{(n)} = n^{(0)} + n, \quad y^{(n)} = \frac{n^{(0)}}{n^{(0)} + n} y^{(0)} + \frac{n}{n^{(0)} + n} \cdot \frac{s}{n}$$

$y^{(0)} = E[p]$

What if expert information and data tell different stories?

- ▶ reparametrisation helps to understand effect of prior-data conflict:

$$n^{(0)} = \alpha^{(0)} + \beta^{(0)}, \quad y^{(0)} = \frac{\alpha^{(0)}}{\alpha^{(0)} + \beta^{(0)}}, \quad \text{which are updated as}$$

$$n^{(n)} = n^{(0)} + n, \quad y^{(n)} = \frac{n^{(0)}}{n^{(0)} + n} y^{(0)} + \frac{n}{n^{(0)} + n} \cdot \frac{s}{n}$$

$$y^{(0)} = E[p] \quad y^{(n)} = E[p | s]$$

What if expert information and data tell different stories?

- ▶ reparametrisation helps to understand effect of prior-data conflict:

$$n^{(0)} = \alpha^{(0)} + \beta^{(0)}, \quad y^{(0)} = \frac{\alpha^{(0)}}{\alpha^{(0)} + \beta^{(0)}}, \quad \text{which are updated as}$$

$$n^{(n)} = n^{(0)} + n, \quad y^{(n)} = \frac{n^{(0)}}{n^{(0)} + n} y^{(0)} + \frac{n}{n^{(0)} + n} \cdot \frac{s}{n}$$

$$y^{(0)} = E[p]$$

$$y^{(n)} = E[p | s]$$

ML estimator  $\hat{p}$

What if expert information and data tell different stories?

- ▶ reparametrisation helps to understand effect of prior-data conflict:

$n^{(0)} = \alpha^{(0)} + \beta^{(0)},$     $y^{(0)} = \frac{\alpha^{(0)}}{\alpha^{(0)} + \beta^{(0)}},$    which are updated as

$n^{(n)} = n^{(0)} + n,$     $y^{(n)} = \frac{n^{(0)}}{n^{(0)} + n} y^{(0)} + \frac{n}{n^{(0)} + n} \cdot \frac{s}{n}$

$n^{(0)} = \text{pseudocounts}$     $y^{(0)} = E[p]$     $y^{(n)} = E[p | s]$    ML estimator  $\hat{p}$

What if expert information and data tell different stories?

- ▶ reparametrisation helps to understand effect of prior-data conflict:

$$n^{(0)} = \alpha^{(0)} + \beta^{(0)}, \quad y^{(0)} = \frac{\alpha^{(0)}}{\alpha^{(0)} + \beta^{(0)}}, \quad \text{which are updated as}$$
$$n^{(n)} = n^{(0)} + n, \quad y^{(n)} = \frac{n^{(0)}}{n^{(0)} + n} y^{(0)} + \frac{n}{n^{(0)} + n} \cdot \frac{s}{n}$$

$n^{(0)} = \text{pseudocounts}$     $y^{(0)} = E[p]$     $y^{(n)} = E[p | s]$    ML estimator  $\hat{p}$

$E[p | s] = y^{(n)}$  is a weighted average of  $E[p]$  and  $\hat{p}$ !

What if expert information and data tell different stories?

- ▶ reparametrisation helps to understand effect of prior-data conflict:

$n^{(0)} = \alpha^{(0)} + \beta^{(0)},$  which are updated as

$n^{(n)} = n^{(0)} + n,$

$y^{(0)} = \frac{\alpha^{(0)}}{\alpha^{(0)} + \beta^{(0)}},$

$y^{(n)} = \frac{n^{(0)}}{n^{(0)} + n} y^{(0)} + \frac{n}{n^{(0)} + n} \cdot \frac{s}{n}$

$n^{(0)}$  = pseudocounts     $y^{(0)}$  =  $E[p]$      $y^{(n)}$  =  $E[p | s]$     ML estimator  $\hat{p}$

$E[p | s] = y^{(n)}$  is a weighted average of  $E[p]$  and  $\hat{p}$ !

$\text{Var}[p | s] = \frac{y^{(n)}(1 - y^{(n)})}{n^{(n)} + 1}$  decreases with  $n$ !

Add **imprecision** as new modelling dimension:  
**Sets of priors...**

... model uncertainty in probability statements

Add **imprecision** as new modelling dimension:  
**Sets of priors...**

... model uncertainty in probability statements

## Uncertainty about probability statements

smaller sets = more precise probability statements

### Lottery A

Number of winning tickets:  
exactly known as 5 out of 100

▶  $P(\text{win}) = 5/100$

### Lottery B

Number of winning tickets:  
not exactly known, supposedly  
between 1 and 7 out of 100

▶  $P(\text{win}) = [1/100, 7/100]$



Add **imprecision** as new modelling dimension:  
**Sets of priors...**

- ... model uncertainty in probability statements
- ... allow for partial or vague information on  $p_t^k$ 's

## Add **imprecision** as new modelling dimension: **Sets of priors...**

- ... model uncertainty in probability statements
- ... allow for partial or vague information on  $p_t^k$ 's
- ... highlight prior-data conflict.

## Add **imprecision** as new modelling dimension: **Sets of priors...**

- ... model uncertainty in probability statements
  - ... allow for partial or vague information on  $p_t^k$ 's
  - ... highlight prior-data conflict.
- ▶ Separate uncertainty *within the model* (reliability statements) from uncertainty *about the model* (which parameters).

## Add **imprecision** as new modelling dimension: **Sets of priors...**

- ... model uncertainty in probability statements
  - ... allow for partial or vague information on  $p_t^k$ 's
  - ... highlight prior-data conflict.
- 
- ▶ Separate uncertainty *within the model* (reliability statements) from uncertainty *about the model* (which parameters).
  - ▶ Systematic sensitivity analysis / robust Bayesian approach

## Add **imprecision** as new modelling dimension: **Sets of priors...**

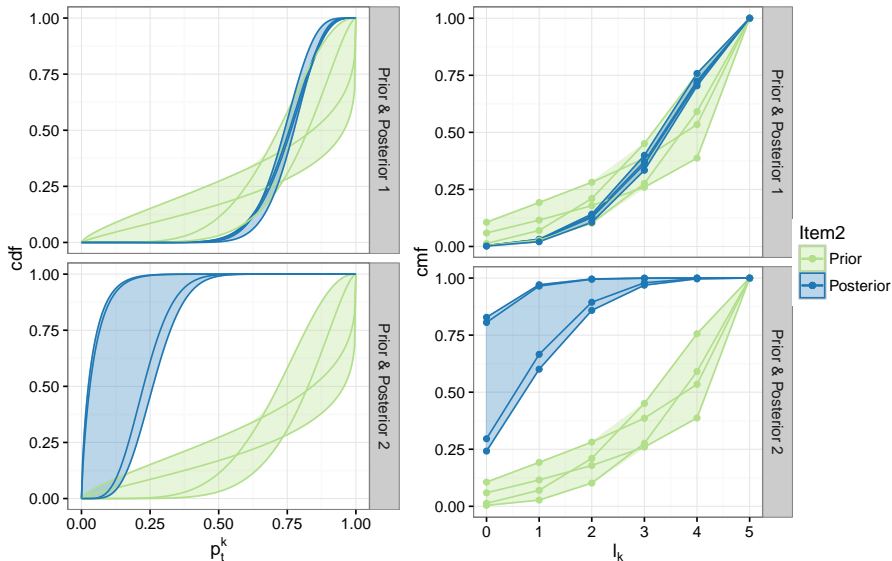
- ... model uncertainty in probability statements
  - ... allow for partial or vague information on  $p_t^k$ 's
  - ... highlight prior-data conflict.
- 
- ▶ Separate uncertainty *within the model* (reliability statements) from uncertainty *about the model* (which parameters).
  - ▶ Systematic sensitivity analysis / robust Bayesian approach
  - ▶ Walter and Augustin (2009), Walter (2013):  
vary  $(\underline{n}^{(0)}, \underline{y}^{(0)})$  in a set  $= [\underline{n}^{(0)}, \bar{n}^{(0)}] \times [\underline{y}^{(0)}, \bar{y}^{(0)}]$ 
    - ▶ easy elicitation, tractability & prior-data conflict sensitivity

## Add **imprecision** as new modelling dimension: **Sets of priors...**

- ... model uncertainty in probability statements
  - ... allow for partial or vague information on  $p_t^k$ 's
  - ... highlight prior-data conflict.
- 
- ▶ Separate uncertainty *within the model* (reliability statements) from uncertainty *about the model* (which parameters).
  - ▶ Systematic sensitivity analysis / robust Bayesian approach
  - ▶ Walter and Augustin (2009), Walter (2013):  
vary  $(n^{(0)}, y^{(0)})$  in a set  $= [\underline{n}^{(0)}, \bar{n}^{(0)}] \times [\underline{y}^{(0)}, \bar{y}^{(0)}]$ 
    - ▶ easy elicitation, tractability & prior-data conflict sensitivity
  - ▶ Bounds for inferences (point estimate, prediction, ...)  
by min/max over

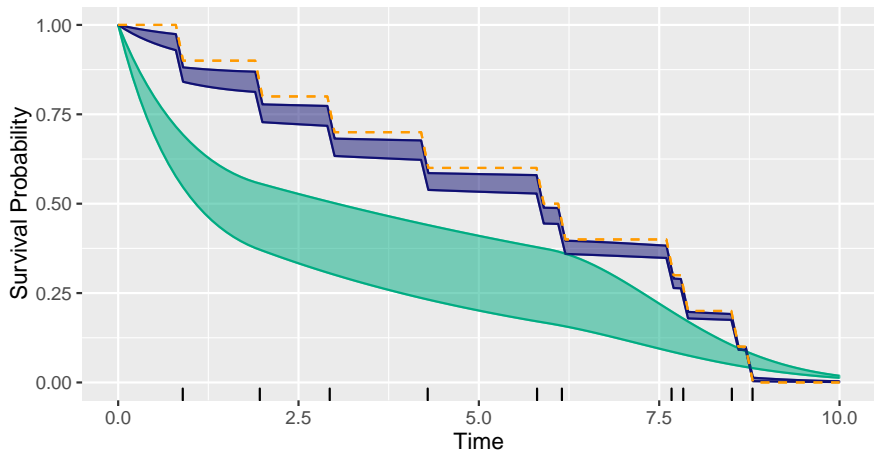
# Sets of Priors for $p_t^k$ and $C_t^k$

6/20



# Component Reliability with Sets of Priors

7/20



$$[\underline{n}^{(0)}, \bar{n}^{(0)}] = [1, 2]$$

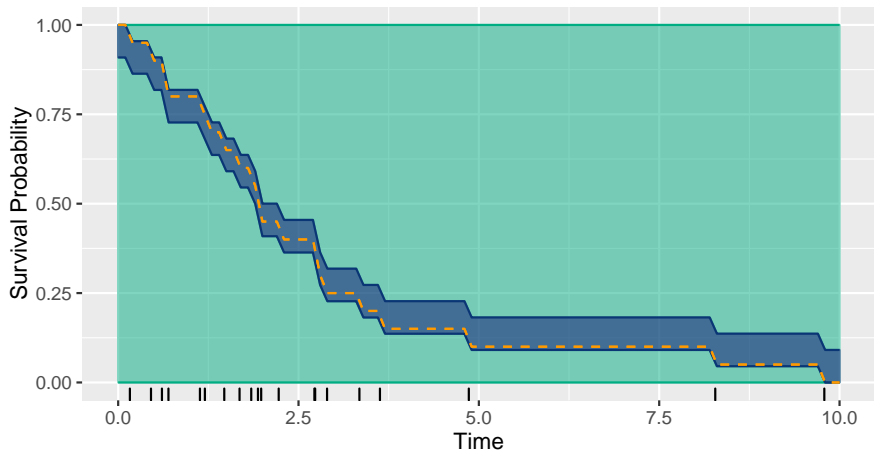
$$[\underline{y}^{(0)}, \bar{y}^{(0)}] = \updownarrow$$

 Prior  Posterior



# Component Reliability with Sets of Priors

7/20



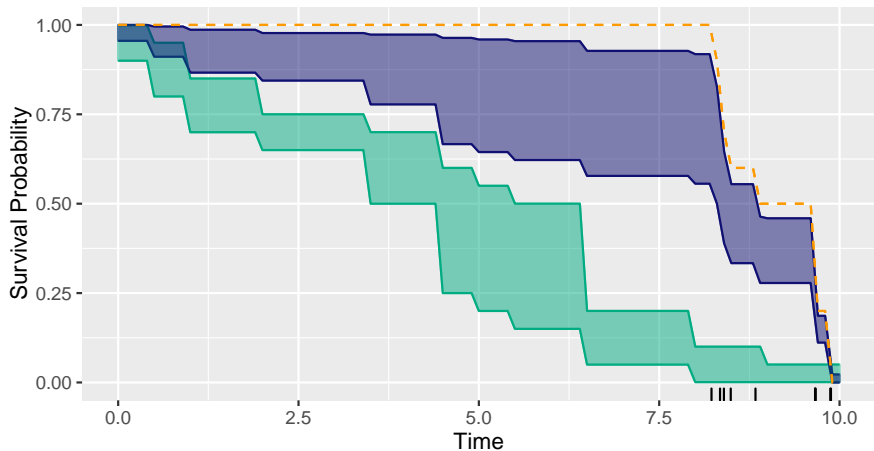
$$[\underline{n}^{(0)}, \bar{n}^{(0)}] = [1, 2]$$

$$[\underline{y}^{(0)}, \bar{y}^{(0)}] = (0, 1)$$

 Prior  Posterior

# Component Reliability with Sets of Priors

7/20



$$[\underline{n}^{(0)}, \bar{n}^{(0)}] = [1, 8]$$

$$[\underline{y}^{(0)}, \bar{y}^{(0)}] = \updownarrow$$

 Prior  Posterior

- ▶ Closed form for the system reliability via the survival signature:

$$\begin{aligned} R_{\text{sys}}(t \mid \{n_{k,t}^{(0)}, y_{k,t}^{(0)}, t^k\}^{k=1:K}) &= P(T_{\text{sys}} > t \mid \dots) \\ &= \sum_{l_1=0}^{m_1} \dots \sum_{l_K=0}^{m_K} \Phi(l_1, \dots, l_K) \prod_{k=1}^K P(C_t^k = l_k \mid n_{k,t}^{(0)}, y_{k,t}^{(0)}, t^k) \end{aligned}$$

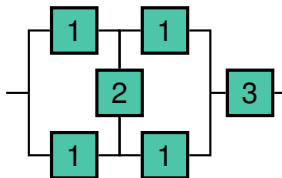
- ▶ Closed form for the system reliability via the survival signature:

$$R_{\text{sys}}(t \mid \{n_{k,t}^{(0)}, y_{k,t}^{(0)}, t^k\}_{k=1:K}) = P(T_{\text{sys}} > t \mid \dots)$$

$$= \sum_{l_1=0}^{m_1} \dots \sum_{l_K=0}^{m_K} \Phi(l_1, \dots, l_K) \prod_{k=1}^K P(C_t^k = l_k \mid n_{k,t}^{(0)}, y_{k,t}^{(0)}, t^k)$$

Survival signature  $\Phi(l_1, \dots, l_K)$   
 (Coolen and Coolen-Maturi 2012)  
 $= P(\text{system functions} \mid \{l_k \text{ k's function}\}_{1:K})$

$l_1$	$l_2$	$l_3$	$\Phi$	$l_1$	$l_2$	$l_3$	$\Phi$
0	0	1	0	0	1	1	0
1	0	1	0	1	1	1	0
2	0	1	1/3	2	1	1	2/3
3	0	1	1	3	1	1	1
4	0	1	1	4	1	1	1



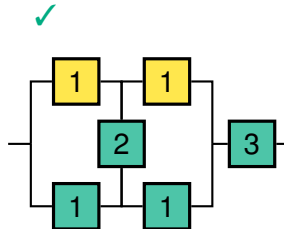
- Closed form for the system reliability via the survival signature:

$$R_{\text{sys}}(t \mid \{n_{k,t}^{(0)}, y_{k,t}^{(0)}, t^k\}^{k=1:K}) = P(T_{\text{sys}} > t \mid \dots)$$

$$= \sum_{l_1=0}^{m_1} \dots \sum_{l_K=0}^{m_K} \Phi(l_1, \dots, l_K) \prod_{k=1}^K P(C_t^k = l_k \mid n_{k,t}^{(0)}, y_{k,t}^{(0)}, t^k)$$

Survival signature  $\Phi(l_1, \dots, l_K)$   
 (Coolen and Coolen-Maturi 2012)  
 $= P(\text{system functions} \mid \{l_k \text{ k's function}\}^{1:K})$

$l_1$	$l_2$	$l_3$	$\Phi$	$l_1$	$l_2$	$l_3$	$\Phi$
0	0	1	0	0	1	1	0
1	0	1	0	1	1	1	0
2	0	1	1/3	2	1	1	2/3
3	0	1	1	3	1	1	1
4	0	1	1	4	1	1	1



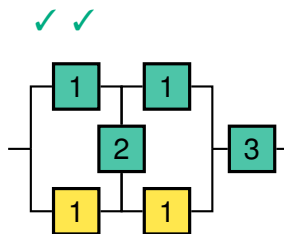
- Closed form for the system reliability via the survival signature:

$$R_{\text{sys}}(t \mid \{n_{k,t}^{(0)}, y_{k,t}^{(0)}, t^k\}^{k=1:K}) = P(T_{\text{sys}} > t \mid \dots)$$

$$= \sum_{l_1=0}^{m_1} \dots \sum_{l_K=0}^{m_K} \Phi(l_1, \dots, l_K) \prod_{k=1}^K P(C_t^k = l_k \mid n_{k,t}^{(0)}, y_{k,t}^{(0)}, t^k)$$

Survival signature  $\Phi(l_1, \dots, l_K)$   
 (Coolen and Coolen-Maturi 2012)  
 $= P(\text{system functions} \mid \{l_k \text{ k's function}\}^{1:K})$

$l_1$	$l_2$	$l_3$	$\Phi$	$l_1$	$l_2$	$l_3$	$\Phi$
0	0	1	0	0	1	1	0
1	0	1	0	1	1	1	0
2	0	1	1/3	2	1	1	2/3
3	0	1	1	3	1	1	1
4	0	1	1	4	1	1	1



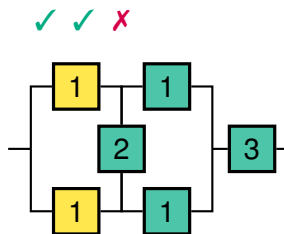
- Closed form for the system reliability via the survival signature:

$$R_{\text{sys}}(t \mid \{n_{k,t}^{(0)}, y_{k,t}^{(0)}, t^k\}^{k=1:K}) = P(T_{\text{sys}} > t \mid \dots)$$

$$= \sum_{l_1=0}^{m_1} \dots \sum_{l_K=0}^{m_K} \Phi(l_1, \dots, l_K) \prod_{k=1}^K P(C_t^k = l_k \mid n_{k,t}^{(0)}, y_{k,t}^{(0)}, t^k)$$

Survival signature  $\Phi(l_1, \dots, l_K)$   
 (Coolen and Coolen-Maturi 2012)  
 $= P(\text{system functions} \mid \{l_k \text{ k's function}\}^{1:K})$

$l_1$	$l_2$	$l_3$	$\Phi$	$l_1$	$l_2$	$l_3$	$\Phi$
0	0	1	0	0	1	1	0
1	0	1	0	1	1	1	0
2	0	1	1/3	2	1	1	2/3
3	0	1	1	3	1	1	1
4	0	1	1	4	1	1	1



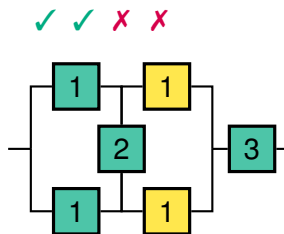
- Closed form for the system reliability via the survival signature:

$$R_{\text{sys}}(t \mid \{n_{k,t}^{(0)}, y_{k,t}^{(0)}, t^k\}^{k=1:K}) = P(T_{\text{sys}} > t \mid \dots)$$

$$= \sum_{l_1=0}^{m_1} \dots \sum_{l_K=0}^{m_K} \Phi(l_1, \dots, l_K) \prod_{k=1}^K P(C_t^k = l_k \mid n_{k,t}^{(0)}, y_{k,t}^{(0)}, t^k)$$

Survival signature  $\Phi(l_1, \dots, l_K)$   
 (Coolen and Coolen-Maturi 2012)  
 $= P(\text{system functions} \mid \{l_k \text{ k's function}\}^{1:K})$

$l_1$	$l_2$	$l_3$	$\Phi$	$l_1$	$l_2$	$l_3$	$\Phi$
0	0	1	0	0	1	1	0
1	0	1	0	1	1	1	0
2	0	1	1/3	2	1	1	2/3
3	0	1	1	3	1	1	1
4	0	1	1	4	1	1	1





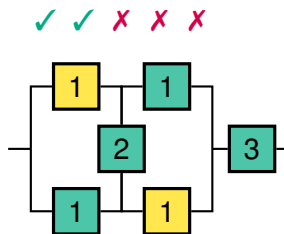
- Closed form for the system reliability via the survival signature:

$$R_{\text{sys}}(t \mid \{n_{k,t}^{(0)}, y_{k,t}^{(0)}, t^k\}_{k=1:K}) = P(T_{\text{sys}} > t \mid \dots)$$

$$= \sum_{l_1=0}^{m_1} \dots \sum_{l_K=0}^{m_K} \Phi(l_1, \dots, l_K) \prod_{k=1}^K P(C_t^k = l_k \mid n_{k,t}^{(0)}, y_{k,t}^{(0)}, t^k)$$

Survival signature  $\Phi(l_1, \dots, l_K)$   
 (Coolen and Coolen-Maturi 2012)  
 $= P(\text{system functions} \mid \{l_k \text{ k's function}\}_{1:K})$

$l_1$	$l_2$	$l_3$	$\Phi$	$l_1$	$l_2$	$l_3$	$\Phi$
0	0	1	0	0	1	1	0
1	0	1	0	1	1	1	0
2	0	1	1/3	2	1	1	2/3
3	0	1	1	3	1	1	1
4	0	1	1	4	1	1	1



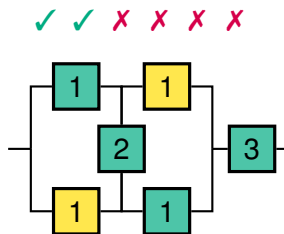
- Closed form for the system reliability via the survival signature:

$$R_{\text{sys}}(t \mid \{n_{k,t}^{(0)}, y_{k,t}^{(0)}, t^k\}_{k=1:K}) = P(T_{\text{sys}} > t \mid \dots)$$

$$= \sum_{l_1=0}^{m_1} \dots \sum_{l_K=0}^{m_K} \Phi(l_1, \dots, l_K) \prod_{k=1}^K P(C_t^k = l_k \mid n_{k,t}^{(0)}, y_{k,t}^{(0)}, t^k)$$

Survival signature  $\Phi(l_1, \dots, l_K)$   
 (Coolen and Coolen-Maturi 2012)  
 $= P(\text{system functions} \mid \{l_k \text{ k's function}\}_{1:K})$

$l_1$	$l_2$	$l_3$	$\Phi$	$l_1$	$l_2$	$l_3$	$\Phi$
0	0	1	0	0	1	1	0
1	0	1	0	1	1	1	0
2	0	1	1/3	2	1	1	2/3
3	0	1	1	3	1	1	1
4	0	1	1	4	1	1	1



- ▶ Closed form for the system reliability via the survival signature:

$$R_{\text{sys}}(t \mid \{n_{k,t}^{(0)}, y_{k,t}^{(0)}, t^k\}^{k=1:K}) = P(T_{\text{sys}} > t \mid \dots)$$

$$= \sum_{l_1=0}^{m_1} \dots \sum_{l_K=0}^{m_K} \Phi(l_1, \dots, l_K) \prod_{k=1}^K P(C_t^k = l_k \mid n_{k,t}^{(0)}, y_{k,t}^{(0)}, t^k)$$

Survival signature  $\Phi(l_1, \dots, l_K)$   
 (Coolen and Coolen-Maturi 2012)  
 $= P(\text{system functions} \mid \{l_k \text{ k's function}\}^{1:K})$

$l_1$	$l_2$	$l_3$	$\Phi$	$l_1$	$l_2$	$l_3$	$\Phi$
0	0	1	0	0	1	1	0
1	0	1	0	1	1	1	0
2	0	1	1/3	2	1	1	2/3
3	0	1	1	3	1	1	1
4	0	1	1	4	1	1	1

Posterior predictive probability that in a new system,  $l_k$  of the  $m_k$  **k**'s function at time  $t$ :

$$\binom{m_k}{l_k} \int_0^1 [p_t^k]^{l_k} [1 - p_t^k]^{m_k - l_k} \times f(p_t^k \mid n_{k,t}^{(0)}, y_{k,t}^{(0)}, t^k) dp_t^k$$

- ▶ analytical solution for integral:  
 $C_t^k \mid n_{k,t}^{(0)}, y_{k,t}^{(0)}, t^k \sim \text{Beta-binomial}$

- ▶ Bounds for  $R_{\text{sys}}(t \mid \{n_{k,t}^{(0)}, y_{k,t}^{(0)}, t^k\}_{k=1:K})$  over  $\mathcal{Y}$ 's:
  - ▶  $\min R_{\text{sys}}(\cdot)$  by  $y_{k,t}^{(0)} = \underline{y}_{k,t}^{(0)}$  for any  $n_{k,t}^{(0)}$   
(Walter, Aslett, and Coolen 2017, Theorem 1)

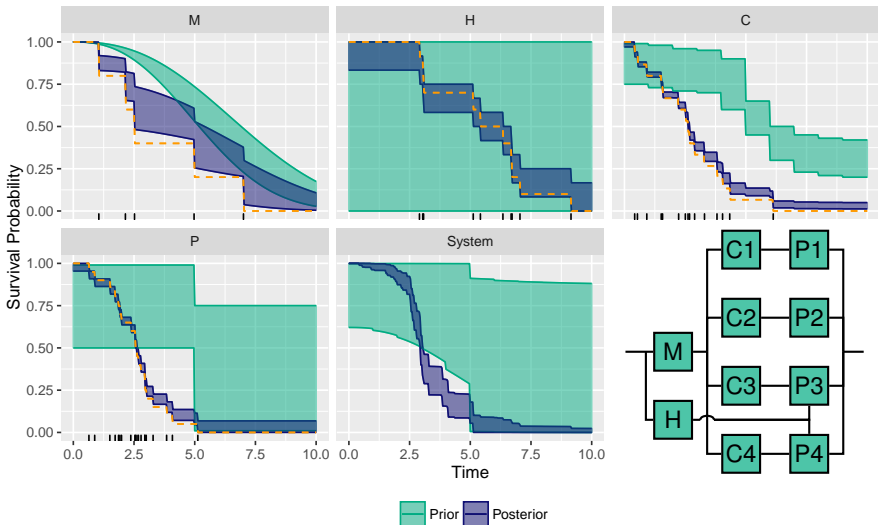
- ▶ Bounds for  $R_{\text{sys}}(t \mid \{n_{k,t}^{(0)}, y_{k,t}^{(0)}, t^k\}_{k=1:K})$  over  $\mathcal{Y}$ 's:
  - ▶  $\min R_{\text{sys}}(\cdot)$  by  $y_{k,t}^{(0)} = \underline{y}_{k,t}^{(0)}$  for any  $n_{k,t}^{(0)}$   
(Walter, Aslett, and Coolen 2017, Theorem 1)
  - ▶  $\min R_{\text{sys}}(\cdot)$  for  $\underline{n}_{k,t}^{(0)}$  or  $\bar{n}_{k,t}^{(0)}$  according to simple conditions  
(Walter, Aslett, and Coolen 2017, Theorem 2 & Lemma 3)

- ▶ Bounds for  $R_{\text{sys}}(t \mid \{n_{k,t}^{(0)}, y_{k,t}^{(0)}, t^k\}_{k=1:K})$  over  $\mathcal{Y}$ 's:
  - ▶  $\min R_{\text{sys}}(\cdot)$  by  $y_{k,t}^{(0)} = \underline{y}_{k,t}^{(0)}$  for any  $n_{k,t}^{(0)}$   
(Walter, Aslett, and Coolen 2017, Theorem 1)
  - ▶  $\min R_{\text{sys}}(\cdot)$  for  $\underline{n}_{k,t}^{(0)}$  or  $\bar{n}_{k,t}^{(0)}$  according to simple conditions  
(Walter, Aslett, and Coolen 2017, Theorem 2 & Lemma 3)
  - ▶ numeric optimization over  $[\underline{n}_{k,t}^{(0)}, \bar{n}_{k,t}^{(0)}]$  in the very few cases where Theorem 2 & Lemma 3 do not apply

- ▶ Bounds for  $R_{\text{sys}}(t \mid \{n_{k,t}^{(0)}, y_{k,t}^{(0)}, t^k\}_{k=1:K})$  over  $\mathbf{y}$ 's:
  - ▶  $\min R_{\text{sys}}(\cdot)$  by  $y_{k,t}^{(0)} = \underline{y}_{k,t}^{(0)}$  for any  $n_{k,t}^{(0)}$   
(Walter, Aslett, and Coolen 2017, Theorem 1)
  - ▶  $\min R_{\text{sys}}(\cdot)$  for  $\underline{n}_{k,t}^{(0)}$  or  $\bar{n}_{k,t}^{(0)}$  according to simple conditions  
(Walter, Aslett, and Coolen 2017, Theorem 2 & Lemma 3)
  - ▶ numeric optimization over  $[\underline{n}_{k,t}^{(0)}, \bar{n}_{k,t}^{(0)}]$  in the very few cases where Theorem 2 & Lemma 3 do not apply
  - ▶ implemented in **R** package `ReliabilityTheory` (Aslett 2016)

# System Reliability Bounds

10/20





## Summary:

- ▶ Nonparametric modeling of component reliability curves
- ▶ Bayesian model combining expert knowledge and test data
- ▶ Set of system reliability functions reflects uncertainties from limited data, vague expert information, and prior-data conflict
- ▶ Easy-to-use implementation in **R** package

ReliabilityTheory (Aslett 2016) with the function  
`nonParBayesSystemInferencePriorSets()`

## Summary:

- ▶ Nonparametric modeling of component reliability curves
- ▶ Bayesian model combining expert knowledge and test data
- ▶ Set of system reliability functions reflects uncertainties from limited data, vague expert information, and prior-data conflict
- ▶ Easy-to-use implementation in **R** package

ReliabilityTheory (Aslett 2016) with the function  
`nonParBayesSystemInferencePriorSets()`

## Next steps:

- ▶ Allow right-censored observations (component monitoring)
- ▶ Allow dependence between components (common-cause failure, ...)
- ▶ Use for system design (where to put extra redundancy?)
- ▶ Use for maintenance planning

# Condition-Based Maintenance Policies for Complex Systems using Component Status Monitoring

Gero Walter, Simme Douwe Flapper

Eindhoven University of Technology, Eindhoven, NL

[g.m.walter@tue.nl](mailto:g.m.walter@tue.nl)

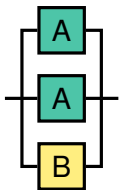


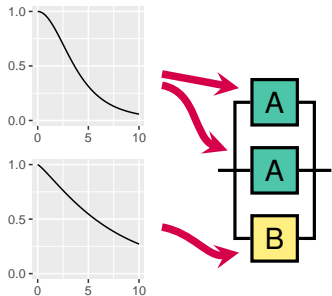
Technische Universiteit  
Eindhoven  
University of Technology

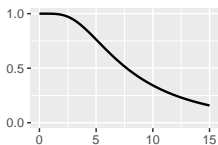
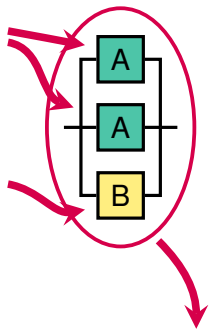
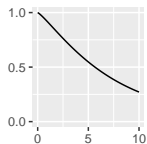
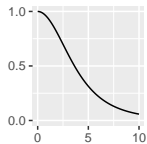


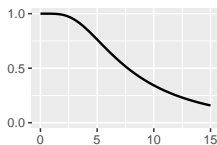
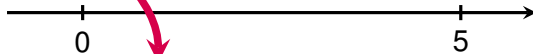
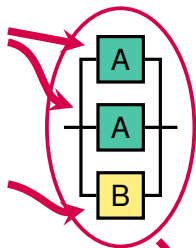
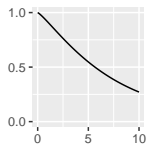
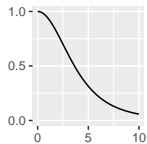
**DINALOG**  
Dutch Institute  
for Advanced Logistics

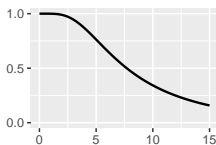
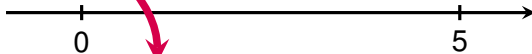
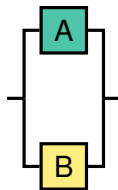
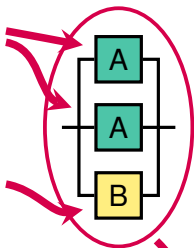
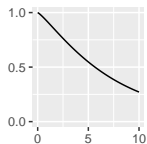
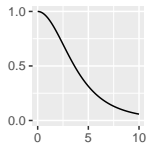
Hannover 2016-12-08



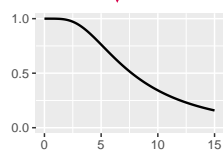
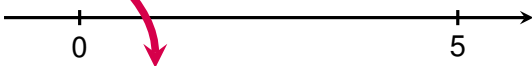
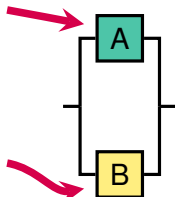
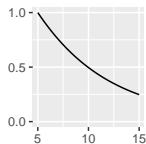
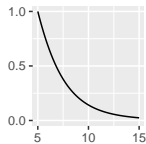
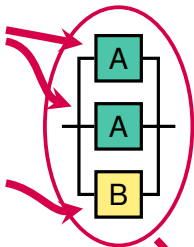
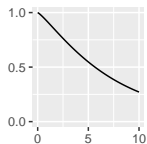
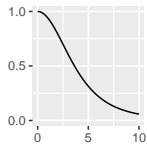


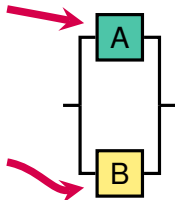
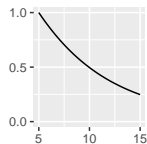
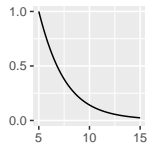
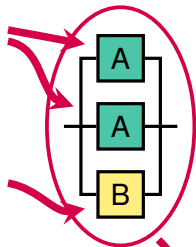
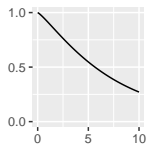
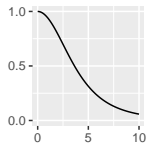




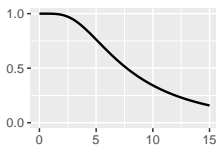
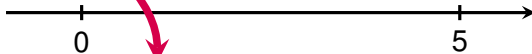


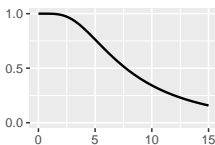
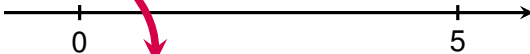
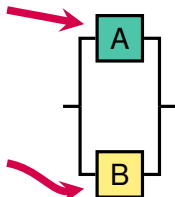
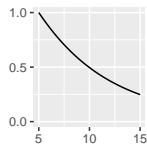
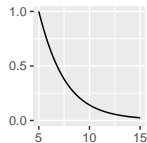
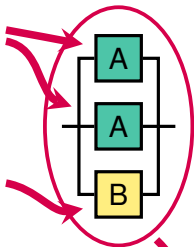
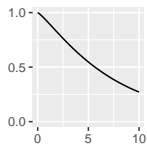
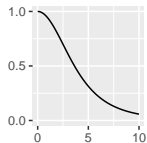




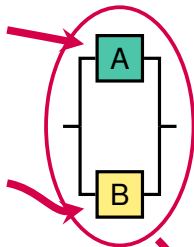
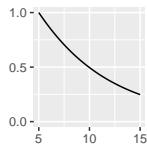
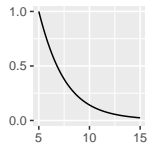
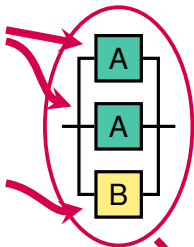
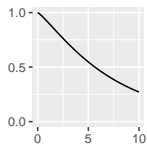
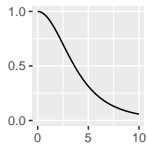


▶ component ageing

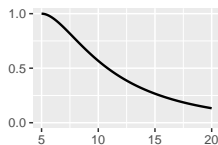
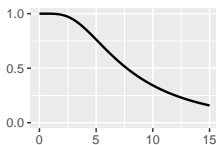
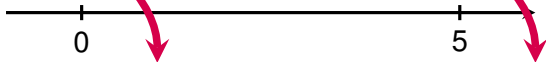


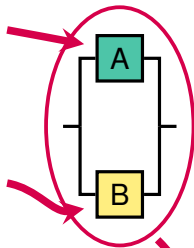
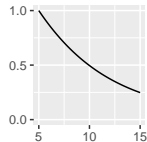
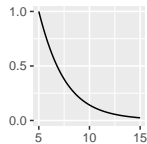
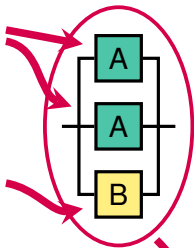
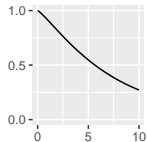
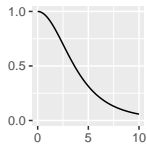


- ▶ component ageing
- ▶ model update

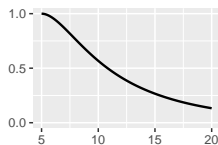
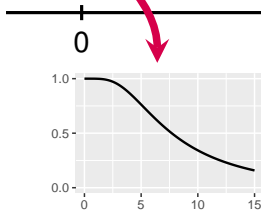


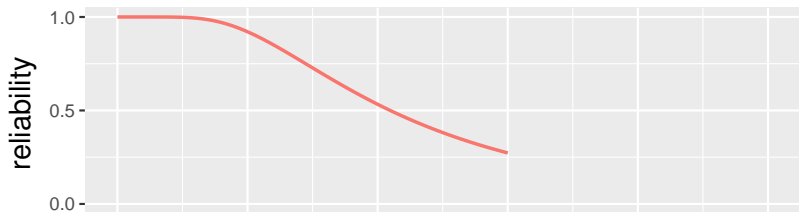
- ▶ component ageing
- ▶ model update

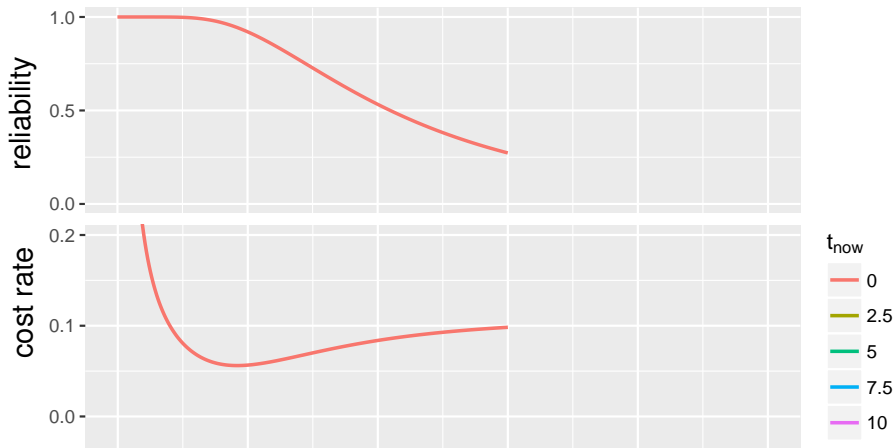


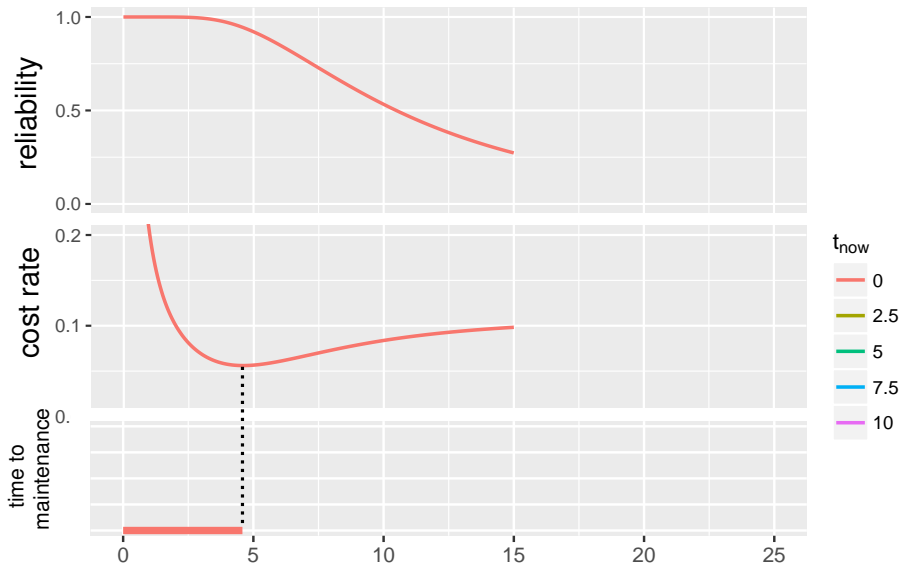


- ▶ component ageing
- ▶ model update
- ▶ system layout

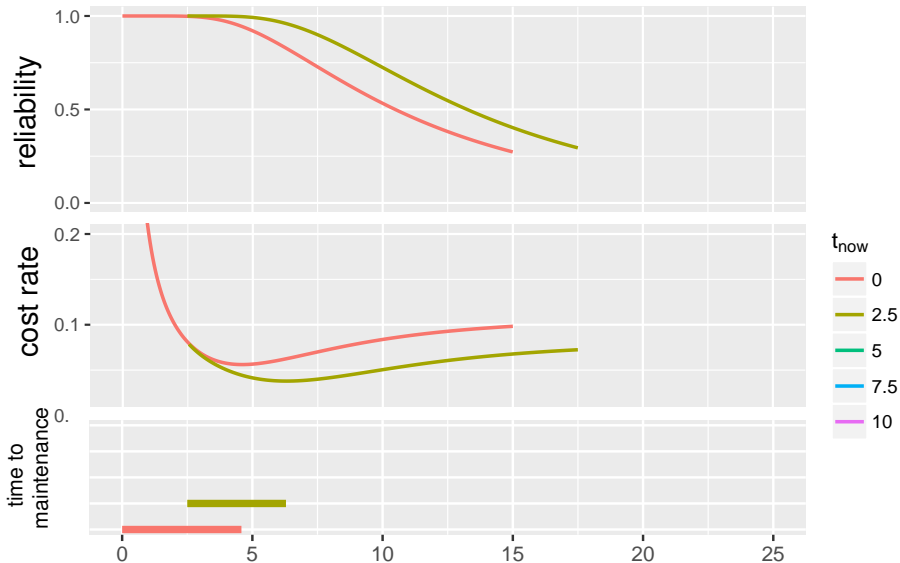


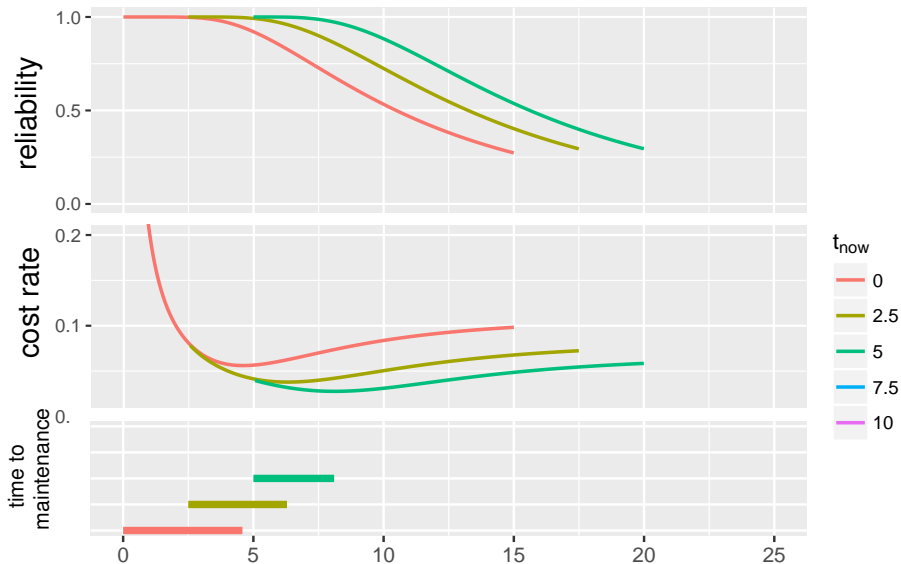


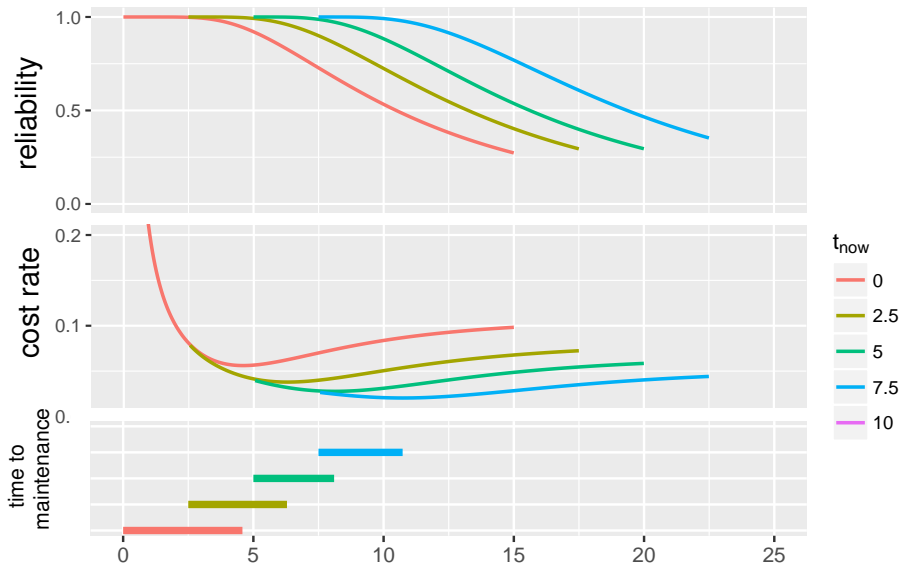


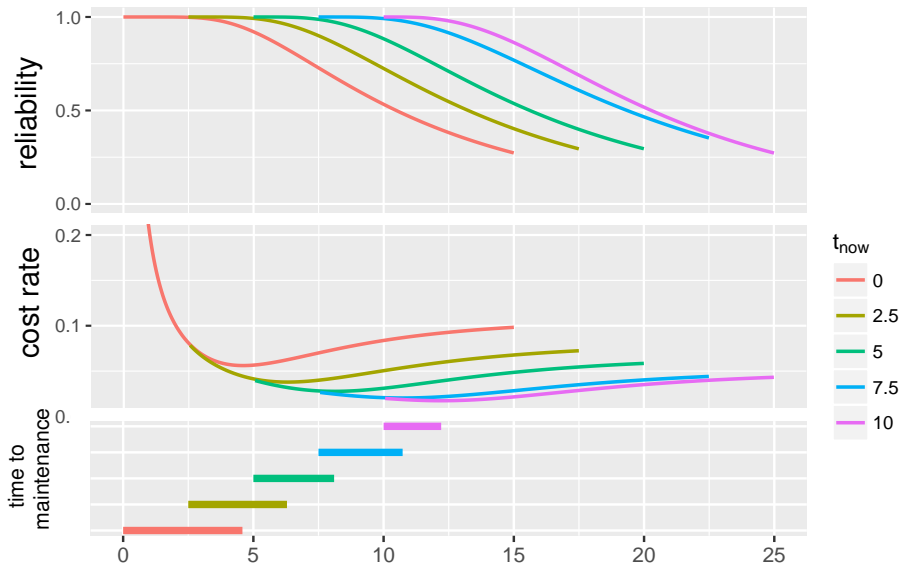


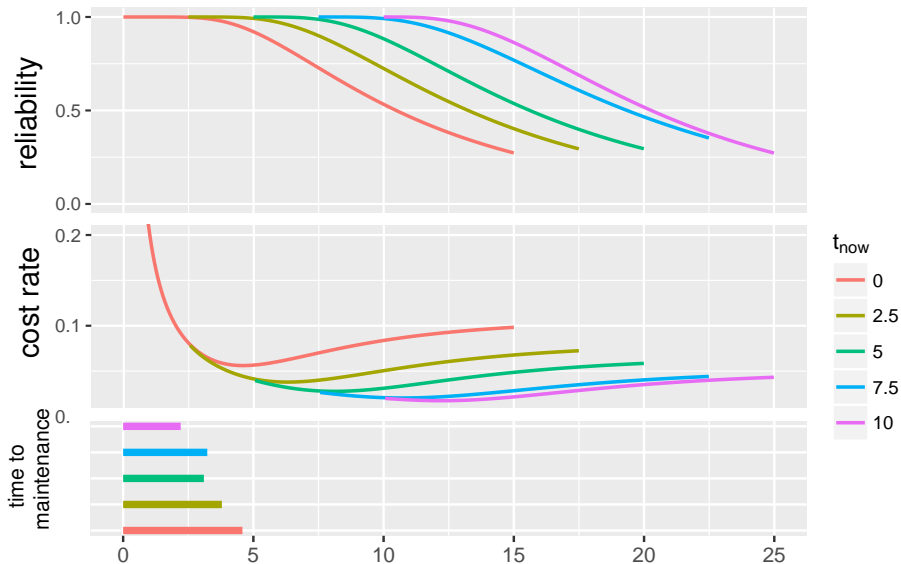


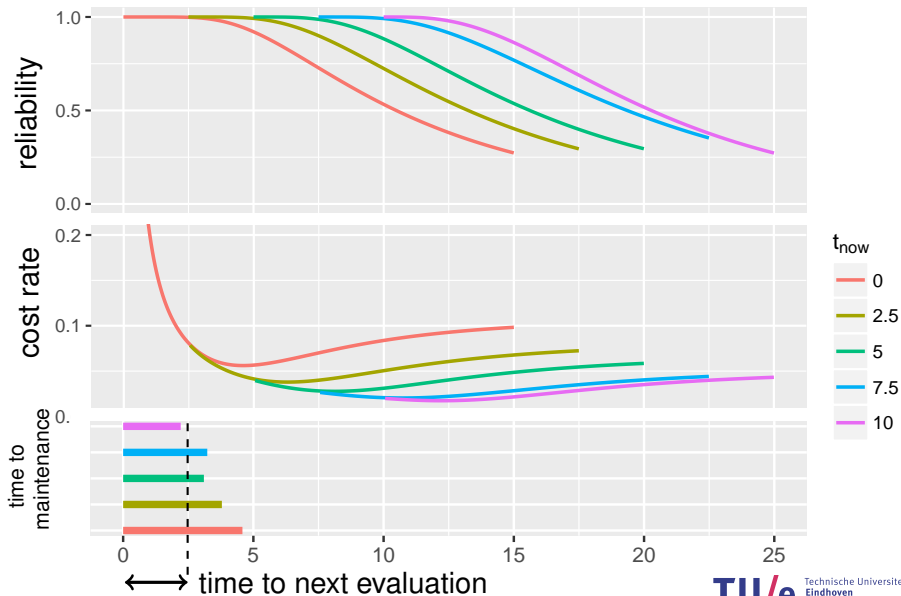


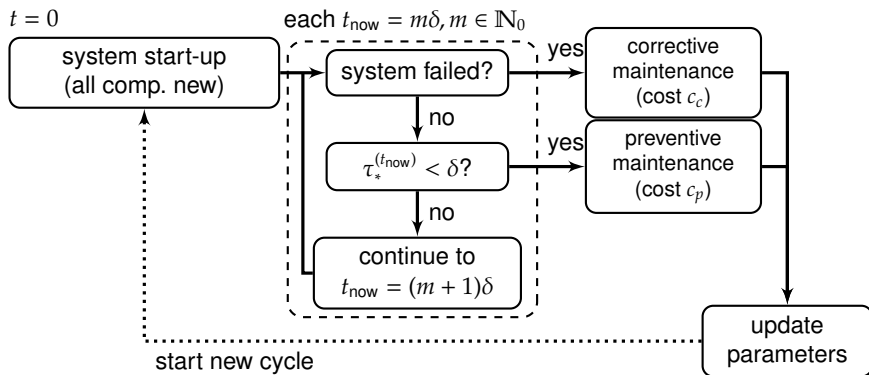


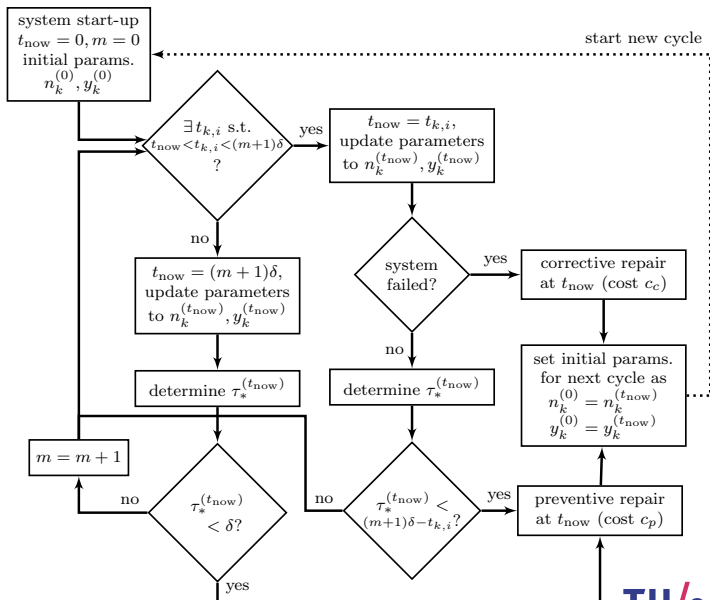














$T_{\text{sys}}^{(t_{\text{now}})}$  (random) time of system failure given all info. at time  $t_{\text{now}}$

$R_{\text{sys}}^{(t_{\text{now}})}(t)$  corresponding reliability function

$c_k^{(t_{\text{now}})}$  number of type  $k$  components functioning at time  $t_{\text{now}}$

$K$  number of component types

$$R_{\text{sys}}^{(t_{\text{now}})}(t) = \sum_{l_1=0}^{c_1^{(t_{\text{now}})}} \cdots \sum_{l_K=0}^{c_K^{(t_{\text{now}})}} \Phi^{(t_{\text{now}})}(l_1, \dots, l_K) \prod_{k=1}^K P(C_t^k = l_k \mid n_k^{(0)}, y_k^{(0)}, \mathbf{t}_k^{(t_{\text{now}})})$$

$T_{\text{sys}}^{(t_{\text{now}})}$  (random) time of system failure given all info. at time  $t_{\text{now}}$

$R_{\text{sys}}^{(t_{\text{now}})}(t)$  corresponding reliability function

$c_k^{(t_{\text{now}})}$  number of type  $k$  components functioning at time  $t_{\text{now}}$

$K$  number of component types

$$R_{\text{sys}}^{(t_{\text{now}})}(t) = \sum_{l_1=0}^{c_1^{(t_{\text{now}})}} \cdots \sum_{l_K=0}^{c_K^{(t_{\text{now}})}} \underbrace{\Phi^{(t_{\text{now}})}(l_1, \dots, l_K)}_{\text{survival signature at time } t_{\text{now}}} \prod_{k=1}^K P(C_t^k = l_k \mid n_k^{(0)}, y_k^{(0)}, t_k^{(t_{\text{now}})})$$

survival signature at time  $t_{\text{now}}$   
=  $P(\text{system functions} \mid \{l_k \mathbf{k}'\text{'s function}\}^{1:K})$

$T_{\text{sys}}^{(t_{\text{now}})}$  (random) time of system failure given all info. at time  $t_{\text{now}}$

$R_{\text{sys}}^{(t_{\text{now}})}(t)$  corresponding reliability function

$c_k^{(t_{\text{now}})}$  number of type  $k$  components functioning at time  $t_{\text{now}}$

$K$  number of component types

$$R_{\text{sys}}^{(t_{\text{now}})}(t) = \sum_{l_1=0}^{c_1^{(t_{\text{now}})}} \cdots \sum_{l_K=0}^{c_K^{(t_{\text{now}})}} \underbrace{\Phi^{(t_{\text{now}})}(l_1, \dots, l_K)}_{\text{survival signature at time } t_{\text{now}}} \prod_{k=1}^K \underbrace{P(C_t^k = l_k \mid n_k^{(0)}, y_k^{(0)}, t_k^{(t_{\text{now}})})}_{\text{Probability that } l_k \text{ of the } c_k^{(t_{\text{now}})} \text{ k's function}}$$

survival signature at time  $t_{\text{now}}$   
 $= P(\text{system functions} \mid \{l_k \text{ k's function}\}^{1:K})$

Probability that  $l_k$  of the  
 $c_k^{(t_{\text{now}})}$  k's function

$\tau$  decision variable (when to do maintenance?)

$T_{\text{sys}}^{(t_{\text{now}})}$  (random) time of system failure,  
with density  $f_{\text{sys}}^{(t_{\text{now}})}(t)$  and reliability function  $R_{\text{sys}}^{(t_{\text{now}})}(t)$

$c_p$  cost of preventive maintenance action

$c_c$  cost of corrective maintenance action

$$g(\tau | T_{\text{sys}}^{(t_{\text{now}})} = t_{\text{now}} + t) = \begin{cases} c_c / (t_{\text{now}} + t) & \text{if } t < \tau \quad (\text{failure before } \tau) \\ c_p / (t_{\text{now}} + \tau) & \text{if } t \geq \tau \quad (\text{failure after } \tau) \end{cases}$$

$\tau$  decision variable (when to do maintenance?)

$T_{\text{sys}}^{(t_{\text{now}})}$  (random) time of system failure,  
with density  $f_{\text{sys}}^{(t_{\text{now}})}(t)$  and reliability function  $R_{\text{sys}}^{(t_{\text{now}})}(t)$

$c_p$  cost of preventive maintenance action

$c_c$  cost of corrective maintenance action

$$g(\tau | T_{\text{sys}}^{(t_{\text{now}})} = t_{\text{now}} + t) = \begin{cases} c_c / (t_{\text{now}} + t) & \text{if } t < \tau \quad (\text{failure before } \tau) \\ c_p / (t_{\text{now}} + \tau) & \text{if } t \geq \tau \quad (\text{failure after } \tau) \end{cases}$$

$$\begin{aligned} g^{(t_{\text{now}})}(\tau) &= \text{E} \left[ g(\tau | T_{\text{sys}}^{(t_{\text{now}})}) \right] \\ &= \frac{c_p}{t_{\text{now}} + \tau} R_{\text{sys}}^{(t_{\text{now}})}(t_{\text{now}} + \tau) + c_c \int_0^{\tau} \frac{1}{t_{\text{now}} + t} f_{\text{sys}}^{(t_{\text{now}})}(t_{\text{now}} + t) dt \end{aligned}$$

$\tau$  decision variable (when to do maintenance?)

$T_{\text{sys}}^{(t_{\text{now}})}$  (random) time of system failure,  
with density  $f_{\text{sys}}^{(t_{\text{now}})}(t)$  and reliability function  $R_{\text{sys}}^{(t_{\text{now}})}(t)$

$c_p$  cost of preventive maintenance action

$c_c$  cost of corrective maintenance action

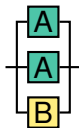
$$g(\tau | T_{\text{sys}}^{(t_{\text{now}})} = t_{\text{now}} + t) = \begin{cases} c_c / (t_{\text{now}} + t) & \text{if } t < \tau \quad (\text{failure before } \tau) \\ c_p / (t_{\text{now}} + \tau) & \text{if } t \geq \tau \quad (\text{failure after } \tau) \end{cases}$$

$$\begin{aligned} g^{(t_{\text{now}})}(\tau) &= \mathbb{E} \left[ g(\tau | T_{\text{sys}}^{(t_{\text{now}})}) \right] \\ &= \frac{c_p}{t_{\text{now}} + \tau} R_{\text{sys}}^{(t_{\text{now}})}(t_{\text{now}} + \tau) + c_c \int_0^{\tau} \frac{1}{t_{\text{now}} + t} f_{\text{sys}}^{(t_{\text{now}})}(t_{\text{now}} + t) dt \end{aligned}$$

$$\tau_*^{(t_{\text{now}})} = \arg \min g^{(t_{\text{now}})}(\tau)$$

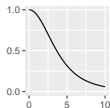
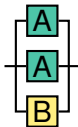
Inputs before start-up:

- ▶ system reliability block diagram



Inputs before start-up:

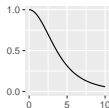
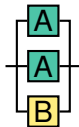
- ▶ system reliability block diagram
- ▶ for each component type:
  - Weibull shape parameter & MTTF from expert
  - expert confidence (how sure about MTTF)
  - optional: test data





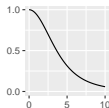
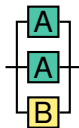
Inputs before start-up:

- ▶ system reliability block diagram
- ▶ for each component type:
  - Weibull shape parameter & MTTF from expert
  - expert confidence (how sure about MTTF)
  - optional: test data
- ▶ cost parameters  $c_p$  and  $c_c$



Inputs before start-up:

- ▶ system reliability block diagram
- ▶ for each component type:
  - Weibull shape parameter & MTTF from expert
  - expert confidence (how sure about MTTF)
  - optional: test data
- ▶ cost parameters  $c_p$  and  $c_c$

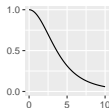
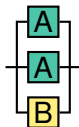


Input during run-time (monitoring):

- ▶ which components still work and which not

Inputs before start-up:

- ▶ system reliability block diagram
- ▶ for each component type:
  - Weibull shape parameter & MTTF from expert
  - expert confidence (how sure about MTTF)
  - optional: test data
- ▶ cost parameters  $c_p$  and  $c_c$



Input during run-time (monitoring):

- ▶ which components still work and which not

Output:

- ▶ for any time during run-time:  
cost-optimal moment to repair the system (dynamic & adaptive)

## Summary:

- ▶ Condition-based maintenance policy for complex system using only component status
- ▶ Weibull component models with known shape parameter & conjugate inverse Gamma prior for scale parameter (no MC or numerical integration for parameter update)
- ▶ minimizes expected cycle cost rate per unit time

## Summary:

- ▶ Condition-based maintenance policy for complex system using only component status
- ▶ Weibull component models with known shape parameter & conjugate inverse Gamma prior for scale parameter (no MC or numerical integration for parameter update)
- ▶ minimizes expected cycle cost rate per unit time

## Outlook:

- ▶ estimate / update also shape parameter (no conjugate prior!)
- ▶ interval-censored failure times, common-cause failures
- ▶ selective component replacement policy
- ▶ sets of inverse Gamma priors / nonparametric component model
  - leads to set of  $R_{\text{sys}}^{(t_{\text{now}})}(t)$  and set of  $g^{(t_{\text{now}})}(t)$
  - what is  $\tau_*^{(t_{\text{now}})}$  then? (needs IP decision criteria)

- Aslett, L. (2016). *ReliabilityTheory: Tools for structural reliability analysis*. R package. URL: <http://www.louisaslett.com>.
- Coolen, F. and T. Coolen-Maturi (2012). “Generalizing the Signature to Systems with Multiple Types of Components”. In: *Complex Systems and Dependability*. Ed. by W. Zamojski et al. Vol. 170. Advances in Intelligent and Soft Computing. Springer, pp. 115–130. DOI: [10.1007/978-3-642-30662-4\\_8](https://doi.org/10.1007/978-3-642-30662-4_8).
- Evans, M. and H. Moshonov (2006). “Checking for Prior-Data Conflict”. In: *Bayesian Analysis* 1, pp. 893–914. URL: <http://projecteuclid.org/euclid.ba/1340370946>.
- Walter, G. (2013). “Generalized Bayesian Inference under Prior-Data Conflict”. PhD thesis. Department of Statistics, LMU Munich. URL: <http://edoc.ub.uni-muenchen.de/17059/>.
- Walter, G., L. Aslett, and F. Coolen (2017). “Bayesian Nonparametric System Reliability using Sets of Priors”. In: *International Journal of Approximate Reasoning* 80, pp. 67–88. DOI: [10.1016/j.ijar.2016.08.005](https://doi.org/10.1016/j.ijar.2016.08.005).
- Walter, G. and T. Augustin (2009). “Imprecision and Prior-data Conflict in Generalized Bayesian Inference”. In: *Journal of Statistical Theory and Practice* 3, pp. 255–271. DOI: [10.1080/15598608.2009.10411924](https://doi.org/10.1080/15598608.2009.10411924).