

Imprecise Probability: General Ideas and Statistical Approaches

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Operational

How can uncertainty be reliably

- ▶ measured?
- ▶ communicated?

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Inference

How can we use our uncertainty model for

- ▶ statistical reasoning?
- ▶ decision making?

Definition

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Notation: A, B, C, \dots

Examples

- ▶ tomorrow, it will rain
- ▶ in the next year, at most 3 components will fail

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how to express our uncertainty regarding events?

Definition

The **probability of an event** is a number between 0 and 1.

Notation: $P(A)$, $P(B)$, $P(C)$, ...

Examples

- ▶ for $A =$ 'tomorrow, it will rain'
my probability $P(A)$ is 0.2
- ▶ for $B =$ 'in the next year, at most 3 components will fail'
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what do these numbers actually mean?
how would you measure them?

Interpretation: Trivial Cases

$P(A) = 0 \iff A$ is practically impossible

logically?

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Interpretation: General Case

- ▶ it's (like) a **frequency**
- ▶ it's a degree of belief (▶ **betting rate**)
- ▶ it's something else

$P(A) = 0.2$ means:

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Frequency Interpretation

- + intuitive, easy to understand
- needs **reference class**, only for **repeatable events**
- needs plenty of data, or strong symmetry assumptions
- ! aleatory

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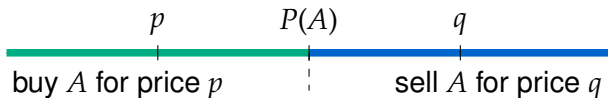
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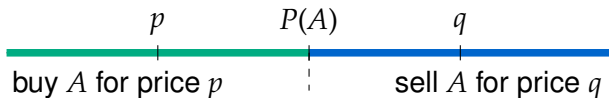
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Betting Interpretation (degree of belief)

- + no reference class, works also for **one-shot events**
- needs plenty of elicitation or plenty of data
- ! epistemic

in case of **partial elicitation** and/or **sparse data**
it may be hard to specify an exact probability
but you may still confidently bound your probability

this becomes more and more relevant
as problems become larger and larger

Confidence intervals (Frequentist Statistics)

- choice of confidence level α ?
- p-value fallacy (Gigerenzer, Krauss, and Vitouch 2004)
a.k.a. prosecutor's fallacy
- + no prior needed, only likelihood

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Interval probability (bounding probabilities directly)

- choice of prior bounds?
- + no confidence / credible level issues
- + no prior ignorance issues
- + no p-value fallacy

Definition

The **lower and upper probability** of an event are numbers between 0 and 1.

Notation: $\underline{P}(A)$, $\overline{P}(A)$, ...

Examples

- ▶ for $A =$ 'tomorrow, it will rain'
my lower probability $\underline{P}(A)$ is 0.1
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$\underline{P}(A) = 0.1$ and $\overline{P}(A) = 0.4$ means:

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- + no reference class, works also for **one-shot events**
- + works with partial elicitation and / or sparse data
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Definition

The **possibility space** Ω is the set of all possible outcomes of the problem at hand.

Example

interested in reliability of a system with 5 components
e.g. number of components that fail in the next year
 $\Omega = \{0, 1, 2, 3, 4, 5\}$

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Definition

An **event** is a subset of Ω . Notation: A, B, C, \dots

Example

in the next year, at most 3 components will fail
would be represented by the event $A = \{0, 1, 2, 3\}$

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A **lower probability** \underline{P} maps every event $A \subseteq \Omega$ to a real number $\underline{P}(A)$.

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- ▶ \underline{P} specification for related events may allow to raise $\underline{P}(A)$ (correcting for consistency)

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maps every event $A \subseteq \Omega$ to a number $P(A)$ in $[0, 1]$ and satisfies

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One of the probability measures P in the credal set \mathcal{M} is the correct one, *but we do not know which*.

crucial: no distribution over \mathcal{M} assumed!

Uncertainty about probability statements

smaller credal set = more precise probability statements

Lottery A

Number of winning tickets:
exactly known as 5 out of 100

▶ $P(\text{win}) = 5/100$

Lottery B

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- ▶ Systematic sensitivity analysis, robust Bayesian approach

expert info + data → complete picture

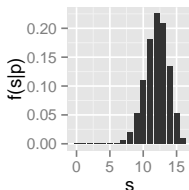
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prior distribution	+	sample distribution	→	posterior distribution
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Binomial
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$$s | p \sim \text{Binomial}(n, p)$$



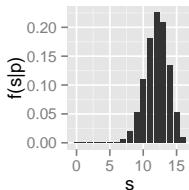
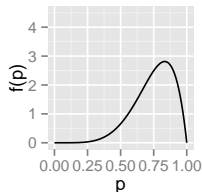
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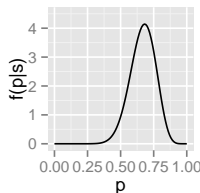
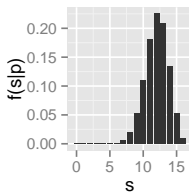
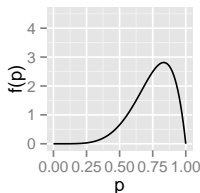
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Beta prior		Binomial distribution	Beta posterior
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▶ Bayes' Rule
▶ conjugacy



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Use the uniform distribution.

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$$E(V) = E(1/\rho) = \int_{0.5}^1 2/\rho \, d\rho = 2(\ln 1 - \ln 0.5) = 1.39\ell$$

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(Jeffreys prior is transformation-invariant,

but depends on the sample space and can break decision making!)

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- ▶ Set of all distributions over volume $V \implies E(V) \in [1, 2]$.
- ▶ Set of all distributions over density $\rho = 1/V$
 $\implies E(V) = E(1/\rho) \in [1, 2]$

Theorem

The set of posterior distributions resulting from a vacuous set of prior distributions is again vacuous, regardless of the likelihood.

We can never learn anything when starting from a vacuous set of priors!

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Solution: Near-Vacuous Sets of Priors

Only insist that the prior predictive, or other classes of inferences, are vacuous.

This can be done using sets of conjugate priors (Walley 1996; Benavoli and Zaffalon 2012; 2015).

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data :	$s p$	\sim	Binomial(n, p)
conjugate prior:	$p \alpha^{(0)}, \beta^{(0)}$	\sim	Beta($\alpha^{(0)}, \beta^{(0)}$)
posterior:	$p \alpha^{(n)}, \beta^{(n)}$	\sim	Beta($\alpha^{(n)}, \beta^{(n)}$)

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Vary hyperparameters ($n^{(0)}, y^{(0)}$) in a set

▶ set of priors $\mathcal{M}^{(0)}$

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Vary hyperparameters $(n^{(0)}, y^{(0)})$ in a set $\mathcal{M}^{(0)}$ ▶ set of priors $\mathcal{M}^{(0)}$

Set of posteriors $\mathcal{M}^{(n)}$ via
$$= \left\{ (n^{(n)}, y^{(n)}) : (n^{(0)}, y^{(0)}) \in \mathcal{M}^{(0)} \right\}$$

Bounds for inferences (point estimate, ...) by min/max over $\mathcal{M}^{(n)}$.

- ▶ reparametrisation helps to understand the parameter update:

$$n^{(0)} = \alpha^{(0)} + \beta^{(0)}, \quad y^{(0)} = \frac{\alpha^{(0)}}{\alpha^{(0)} + \beta^{(0)}}, \quad \text{which are updated as}$$

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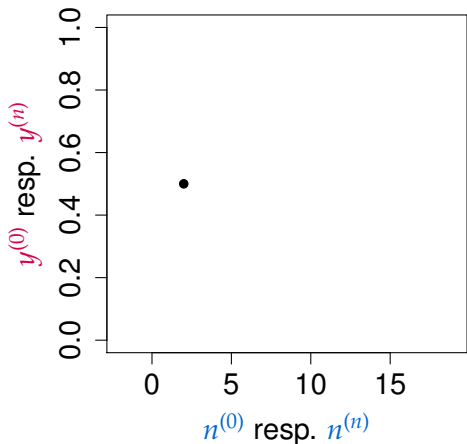
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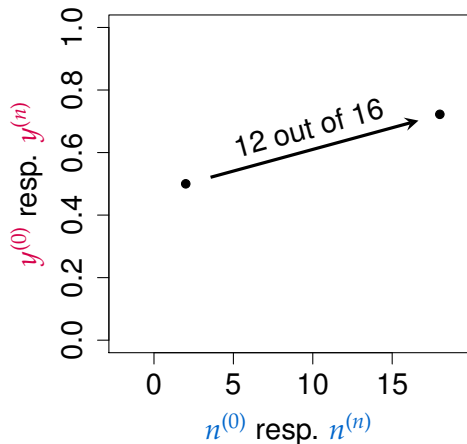
$\text{Var}[p | s] = \frac{y^{(n)}(1 - y^{(n)})}{n^{(n)} + 1}$ decreases with n !



single prior (uniform)

prior $n^{(0)} = 2$, $y^{(0)} = 0.5$

data $s/n = 12/16 = 0.75$

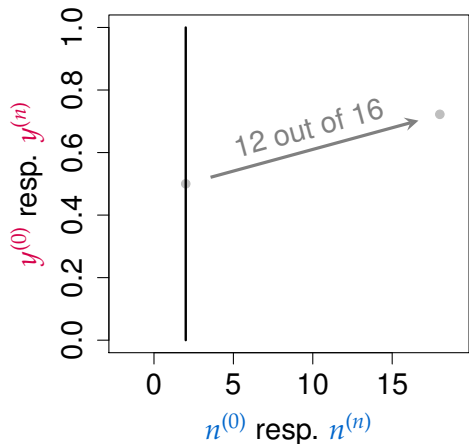


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$n^{(n)} = 18$, $y^{(n)} = 0.72$



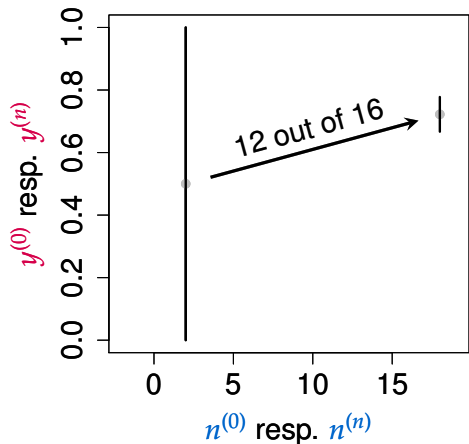
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near-vacuous set of priors

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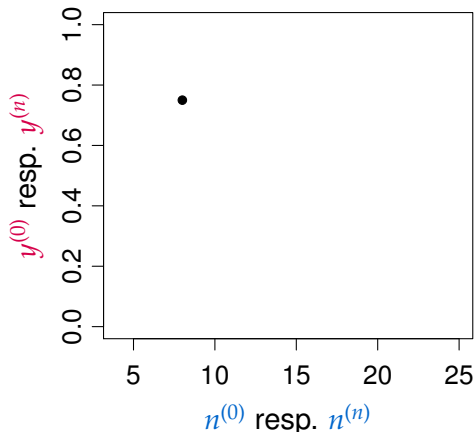
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What if expert information and data tell different stories?

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Prior-Data Conflict

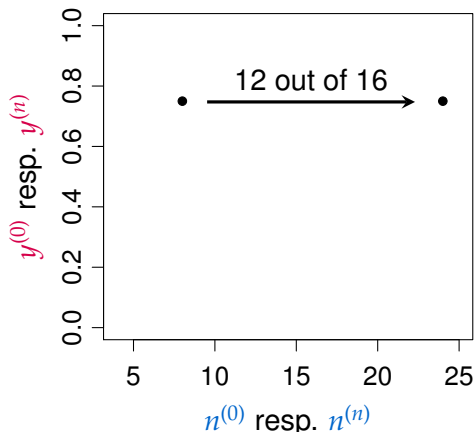
- ▶ *informative prior beliefs* and *trusted data* (sampling model correct, no outliers, etc.) are in conflict
- ▶ “[...] the prior [places] its mass primarily on distributions in the sampling model for which the observed data is surprising” (Evans and Moshonov 2006)
- ▶ there are not enough data to overrule the prior



no conflict:

prior $n^{(0)} = 8$, $y^{(0)} = 0.75$

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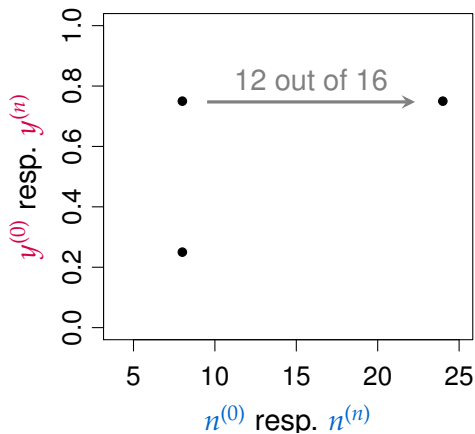
no conflict:

prior $n^{(0)} = 8$, $y^{(0)} = 0.75$

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$n^{(n)} = 24$, $y^{(n)} = 0.75$



no conflict:

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data $s/n = 12/16 = 0.75$

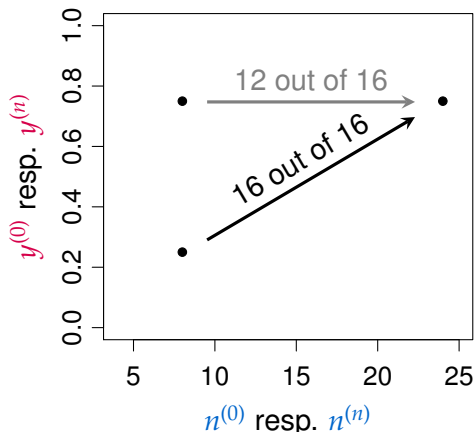


$n^{(n)} = 24$, $y^{(n)} = 0.75$

prior-data conflict:

prior $n^{(0)} = 8$, $y^{(0)} = 0.25$

data $s/n = 16/16 = 1$



no conflict:

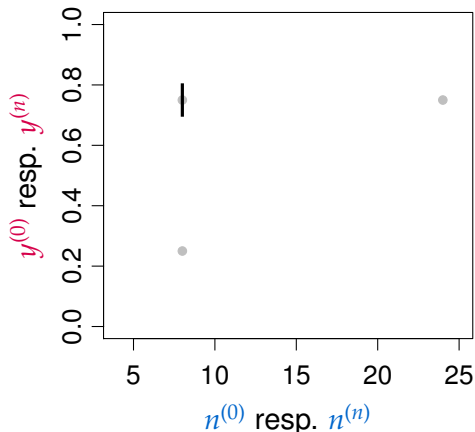
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prior-data conflict:

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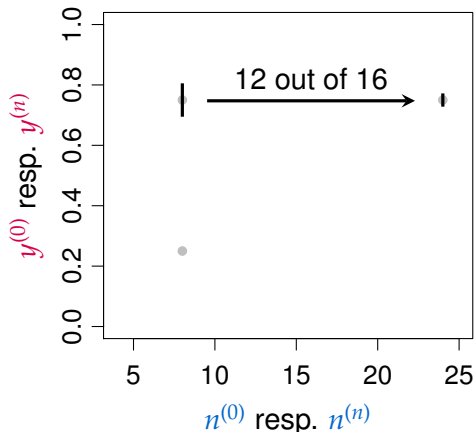
IDM (Walley 1996); Quaeghebeur and de Cooman (2005)



no conflict:

prior $n^{(0)} = 8$, $y^{(0)} \in [0.7, 0.8]$
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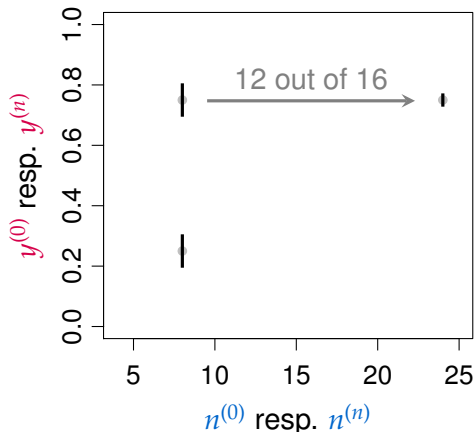


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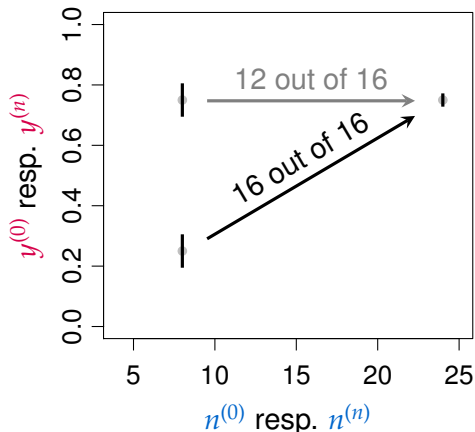
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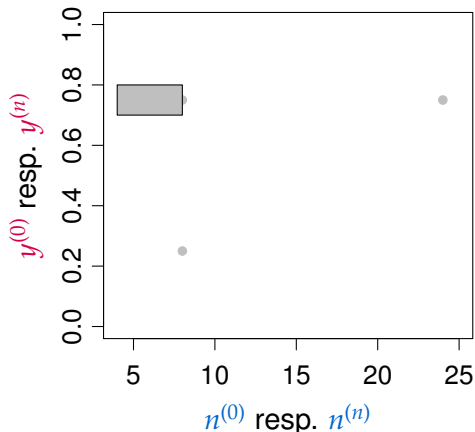
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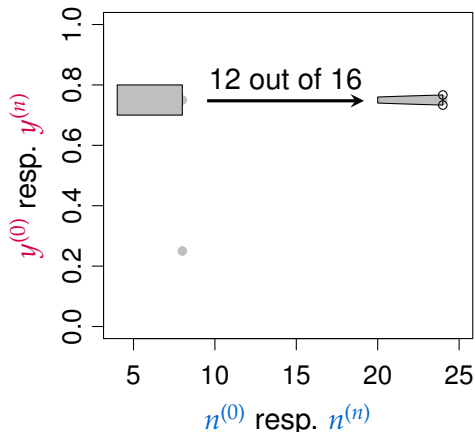


no conflict:

prior $n^{(0)} \in [4, 8]$, $y^{(0)} \in [0.7, 0.8]$

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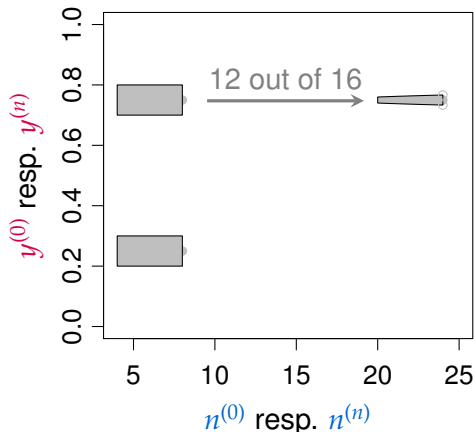


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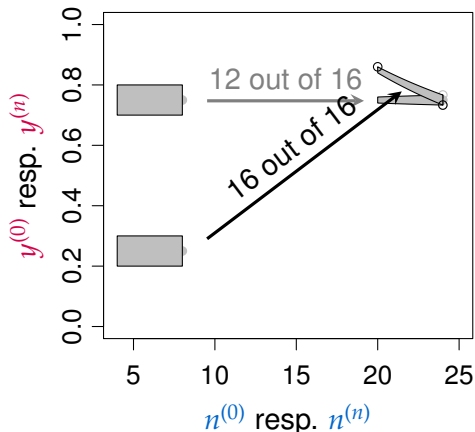
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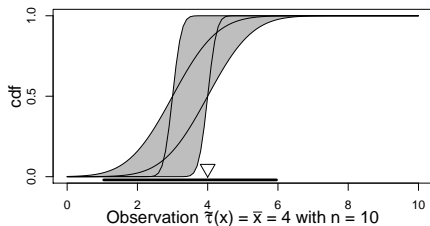
Example: Scaled Normal Data

Data :	$x \mid \mu$	\sim	$N(\mu, 1)$
conjugate prior:	$\mu \mid n^{(0)}, y^{(0)}$	\sim	$N(y^{(0)}, 1/n^{(0)})$
posterior:	$\mu \mid n^{(n)}, y^{(n)}$	\sim	$N(y^{(n)}, 1/n^{(n)}) \quad (\tau(\mathbf{x})/n = \bar{x})$

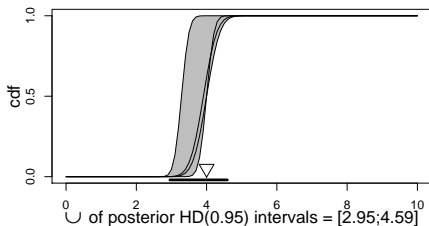
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27/37

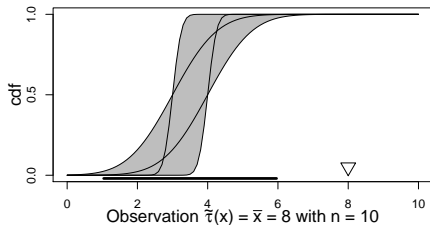
Set of priors: $y^{(0)} \in [3;4]$ and $n^{(0)} \in [1;25]$



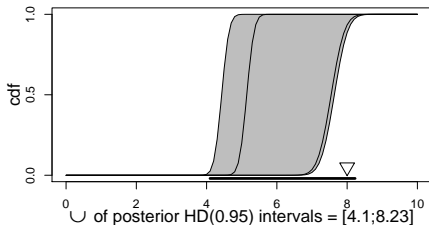
Set of posteriors: $y^{(1)} \in [3.29;4]$ and $n^{(1)} \in [11;35]$



Set of priors: $y^{(0)} \in [3;4]$ and $n^{(0)} \in [1;25]$



Set of posteriors: $y^{(1)} \in [4.43;7.64]$ and $n^{(1)} \in [11;35]$



Conjugate priors like the Beta can be constructed for sample distributions (likelihood) from:

Definition (Canonical exponential family)

$$f(x | \psi) = h(x) \exp \{ \psi^T \tau(x) - b(\psi) \}$$

- ▶ includes multinomial, normal, Poisson, exponential, ...
- ▶ ψ generally a transformation of original parameter θ

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Definition (Family of conjugate priors)

A family of priors for i.i.d. sampling from the can. exp. family:

$$f(\psi | n^{(0)}, y^{(0)}) \propto \exp \{ n^{(0)} [\psi^T y^{(0)} - b(\psi)] \}$$

with hyper-parameters $n^{(0)}$ and $y^{(0)}$.

Theorem (Conjugacy)

Posterior is of the same form as the prior:

$$f(\psi \mid n^{(0)}, \mathbf{y}^{(0)}, \mathbf{x}) \propto \exp \left\{ n^{(n)} \left[\psi^T \mathbf{y}^{(n)} - b(\psi) \right] \right\}$$

where

$$\mathbf{x} = (x_1, \dots, x_n) \quad \tau(\mathbf{x}) = \sum_{i=1}^n \tau(x_i)$$

$$n^{(n)} = n^{(0)} + n \quad \mathbf{y}^{(n)} = \frac{n^{(0)}}{n^{(0)} + n} \cdot \mathbf{y}^{(0)} + \frac{n}{n^{(0)} + n} \cdot \frac{\tau(\mathbf{x})}{n}$$

- ▶ $\mathbf{y}^{(0)}$ = prior expectation of $\tau(\mathbf{x})/n$
- ▶ $n^{(0)}$ determines spread and learning speed

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- ▶ $n \rightarrow \infty$

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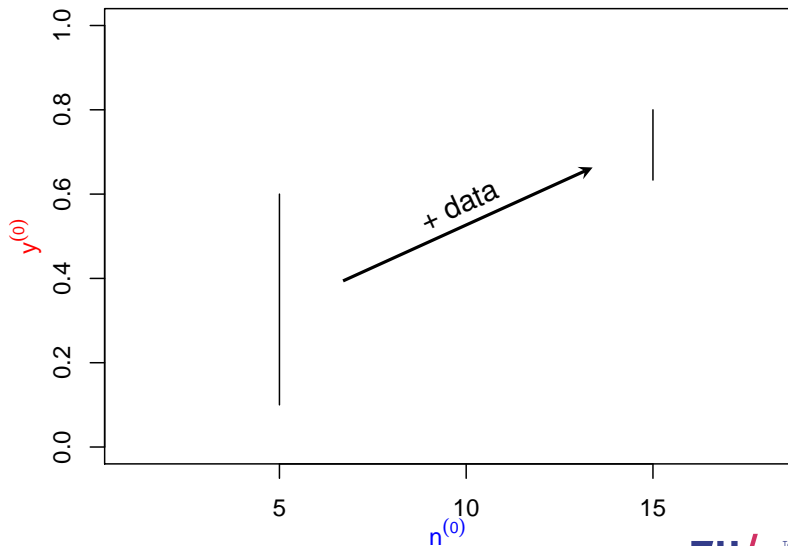
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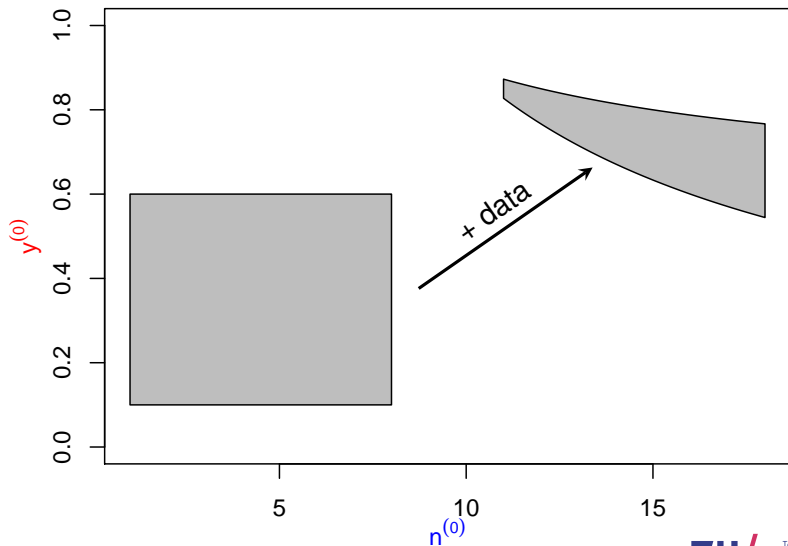
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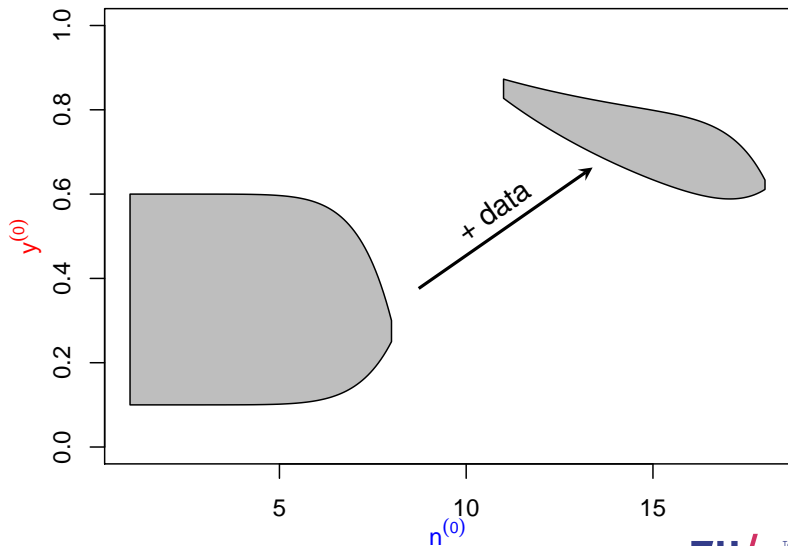
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- ▶ Often, optimising over $(n^{(n)}, y^{(n)}) \in$ is also easy:
closed form solution for $y^{(n)}$ = posterior 'guess' for $\frac{\tau(\mathbf{x})}{n}$ (think: \bar{x})
when has 'nice' shape







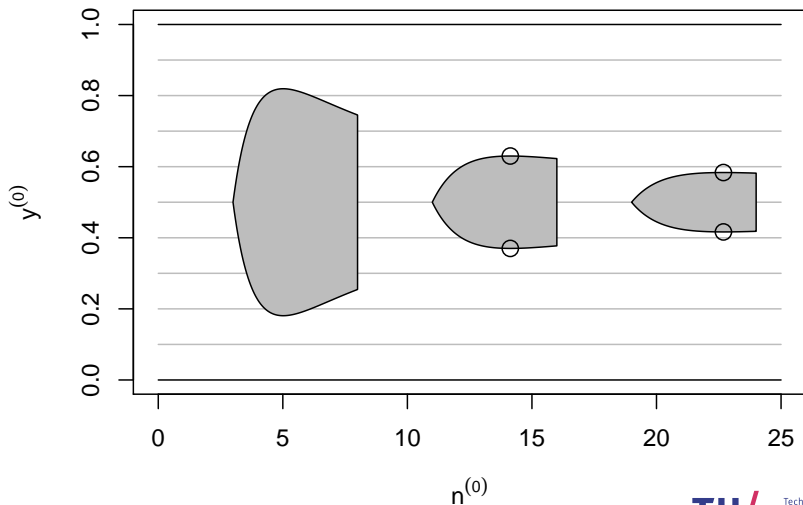
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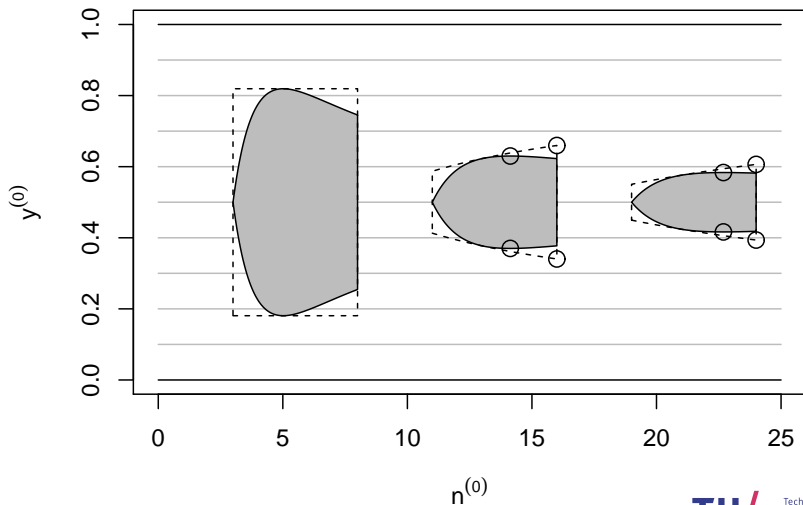
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- ▶ Other set shapes possible, but may be more difficult to handle

set shape for strong prior-data agreement (Walter and Coolen 2016)



set shape for strong prior-data agreement (Walter and Coolen 2016)



- ▶ How to define sets of priors $\mathcal{M}^{(0)}$ is a crucial modeling choice
- ▶ Sets $\mathcal{M}^{(0)}$ via parameter sets $\Pi^{(0)}$ seem to work better than other models discussed in the robust Bayes literature:
 - Neighbourhood models
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 - Density ratio class / interval of measures
 - set of distributions by bounds for the density function $f(\vartheta)$:

$$\mathcal{M}_{l,u}^{(0)} = \{f(\theta) : \exists c \in \mathbb{R}_{>0} : l(\theta) \leq cf(\theta) \leq u(\theta)\}$$

- posterior set is bounded by updated $l(\theta)$ and $u(\theta)$
- $u(\theta)/l(\theta)$ is constant under updating
 - ▶ size of the set does not decrease with n
 - ▶ very vague posterior inferences

- ▶ choice of prior can severely affect inferences even if your prior is 'non-informative'
- ▶ solution: go imprecise
- ▶ models from canonical exponential family make this easy to do (Quaeghebeur and de Cooman 2005; Benavoli and Zaffalon 2012; 2015)
- ▶ allows to adequately express the quality of prior information
- ▶ close relations to robust Bayes literature (Berger et al. 1994; Ríos Insua and Ruggeri 2000)
- ▶ concerns uncertainty in the prior (uncertainty in data generating process: imprecise sampling models)
- ▶ if your prior is informative then prior-data conflict can be an issue (Walter and Augustin 2009; Walter 2013)

- Benavoli, A. and M. Zaffalon (2012). “A model of prior ignorance for inferences in the one-parameter exponential family”. In: *Journal of Statistical Planning and Inference* 142, pp. 1960–1979. DOI: [10.1016/j.jspi.2012.01.023](https://doi.org/10.1016/j.jspi.2012.01.023).
- Benavoli, A. and M. Zaffalon (2015). “Prior near ignorance for inferences in the k-parameter exponential family”. In: *Statistics* 49.5, pp. 1104–1140. DOI: [10.1080/02331888.2014.960869](https://doi.org/10.1080/02331888.2014.960869).
- Berger et al. (1994). “An overview of robust Bayesian analysis”. In: *TEST* 3, pp. 5–124.
- Evans, M. and H. Moshonov (2006). “Checking for Prior-Data Conflict”. In: *Bayesian Analysis* 1, pp. 893–914. URL: <http://projecteuclid.org/euclid.ba/1340370946>.
- Gigerenzer, G., S. Krauss, and O. Vitouch (2004). “The Null Ritual – What you always wanted to know about significance testing but were afraid to ask”. In: *The Sage handbook of quantitative Methodology for the social sciences*. Ed. by D. Kaplan. Thousand Oaks, CA: Sage, pp. 391–408.
- Quaeghebeur, E. and G. de Cooman (2005). “Imprecise probability models for inference in exponential families”. In: *ISIPTA '05. Proceedings of the Fourth International Symposium on Imprecise Probabilities and Their Applications*. Ed. by F. Cozman, R. Nau, and T. Seidenfeld. Manno: SIPTA, pp. 287–296. URL: <http://leo.ugr.es/sipta/isipta05/proceedings/papers/s019.pdf>.
- Ríos Insua, D. and F. Ruggeri (2000). *Robust Bayesian Analysis*. New York: Springer.
- Walley, P. (1991). *Statistical Reasoning with Imprecise Probabilities*. London: Chapman and Hall.

- Walley, P. (1996). "Inferences from multinomial data: Learning about a bag of marbles". In: *Journal of the Royal Statistical Society, Series B* 58.1, pp. 3–34. URL: <http://www.jstor.org/stable/2346164>.
- Walter, G. (2013). "Generalized Bayesian Inference under Prior-Data Conflict". PhD thesis. Department of Statistics, LMU Munich. URL: <http://edoc.ub.uni-muenchen.de/17059/>.
- Walter, G. and T. Augustin (2009). "Imprecision and Prior-data Conflict in Generalized Bayesian Inference". In: *Journal of Statistical Theory and Practice* 3, pp. 255–271. DOI: [10.1080/15598608.2009.10411924](https://doi.org/10.1080/15598608.2009.10411924).
- Walter, G. and F. P. A. Coolen (2016). "Sets of Priors Reflecting Prior-Data Conflict and Agreement". In: *Information Processing and Management of Uncertainty in Knowledge-Based Systems: 16th International Conference, IPMU 2016, Eindhoven, The Netherlands, June 20-24, 2016, Proceedings, Part I*. Ed. by Paulo Joao Carvalho et al. Cham: Springer International Publishing, pp. 153–164. ISBN: 978-3-319-40596-4. DOI: [10.1007/978-3-319-40596-4_14](https://doi.org/10.1007/978-3-319-40596-4_14).
- Walter, G. and N. Krautenbacher (2013). *luck: R package for Generalized iLUCK-models*. URL: <http://luck.r-forge.r-project.org/>.