

# Nonparametric Bayesian System Reliability with Sets of Priors

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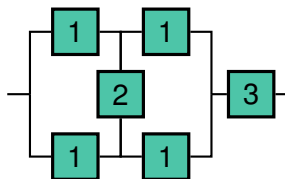
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2016-03-17

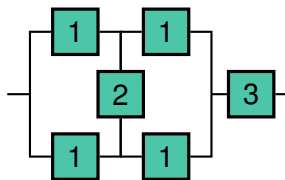
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 $R_{\text{sys}}(t) = P(T_{\text{sys}} > t)$  (system survival function)  
based on

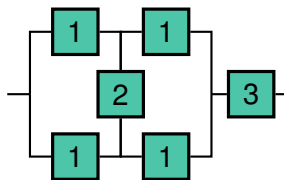


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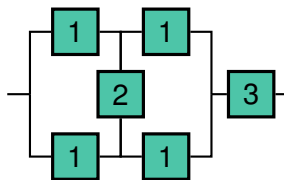
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How to combine these two information sources?

expert info + data → complete picture

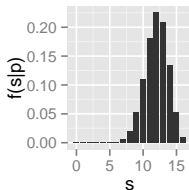
expert info	+	data	→	complete picture
prior distribution	+	sample distribution	→	posterior distribution
$f(p)$	×	$f(s   p)$	∝	$f(p   s)$
				▶ Bayes' Rule

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prior distribution	+	sample distribution	→ posterior distribution
$f(p)$	×	$f(s   p)$	$\propto f(p   s)$

► Bayes' Rule

## Binomial distribution

$$s | p \sim \text{Binomial}(n, p)$$





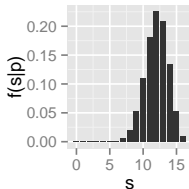
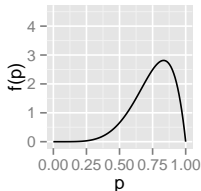
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Beta prior

Binomial  
distribution

$$p \sim \text{Beta}(\alpha^{(0)}, \beta^{(0)})$$

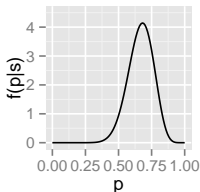
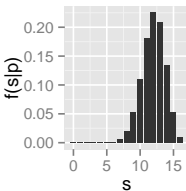
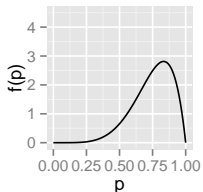
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Beta prior		Binomial distribution	Beta posterior
$p \sim \text{Beta}(\alpha^{(0)}, \beta^{(0)})$		$s   p \sim \text{Binomial}(n, p)$	$p   s \sim \text{Beta}(\alpha^{(n)}, \beta^{(n)})$

► Bayes' Rule

► conjugacy



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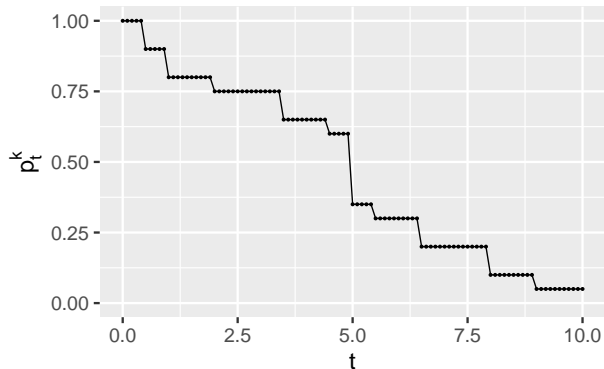
$$s | p \sim \text{Binomial}(n, p)$$

$$p | s \sim \text{Beta}(\alpha^{(n)}, \beta^{(n)})$$

- ▶ conjugate prior makes learning about parameter tractable, just update hyperparameters:  $\alpha^{(0)} \rightarrow \alpha^{(n)}, \beta^{(0)} \rightarrow \beta^{(n)}$
- ▶ closed form for some inferences:  $E[p | s] = \frac{\alpha^{(n)}}{\alpha^{(n)} + \beta^{(n)}}$

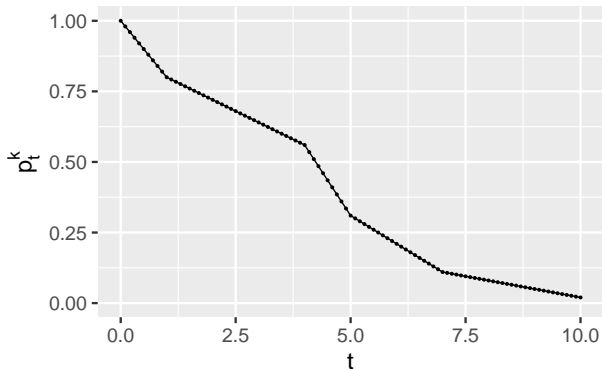
- Functioning probability  $p_t^k$  of **k** for each time  $t \in \mathcal{T} = \{t_1, t_2, \dots\}$
- ▶ discrete component reliability function  $R^k(t) = p_t^k, t \in \mathcal{T}$ .

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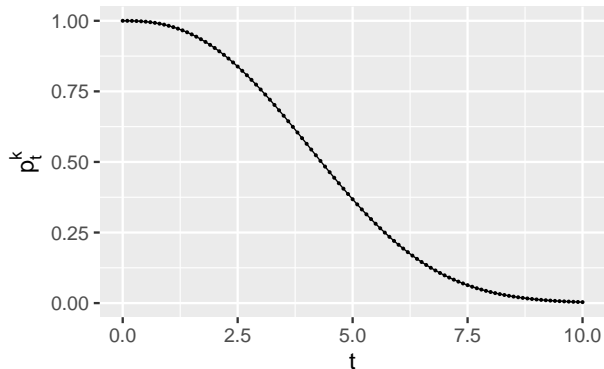
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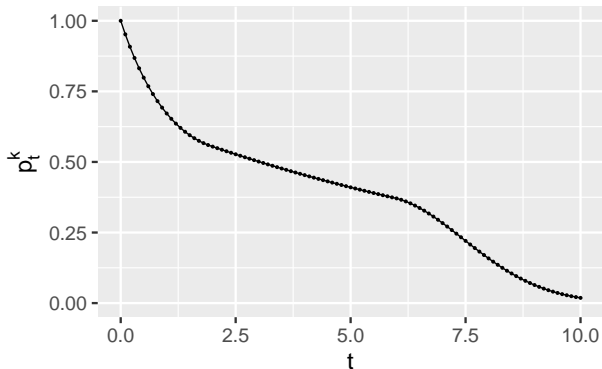
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What if expert information and data tell different stories?

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## Prior-Data Conflict

- ▶ *informative prior beliefs* and *trusted data* (sampling model correct, no outliers, etc.) are in conflict
- ▶ “[. . .] the prior [places] its mass primarily on distributions in the sampling model for which the observed data is surprising” (Evans and Moshonov 2006)
- ▶ there are not enough data to overrule the prior

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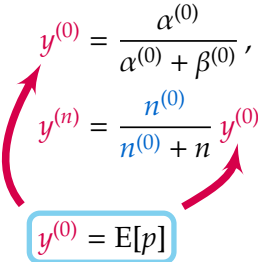
$$n^{(0)} = \alpha^{(0)} + \beta^{(0)}, \quad y^{(0)} = \frac{\alpha^{(0)}}{\alpha^{(0)} + \beta^{(0)}}, \quad \text{which are updated as}$$

$$n^{(n)} = n^{(0)} + n, \quad y^{(n)} = \frac{n^{(0)}}{n^{(0)} + n} y^{(0)} + \frac{n}{n^{(0)} + n} \cdot \frac{s}{n}$$



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$\text{Var}[p | s] = \frac{y^{(n)}(1 - y^{(n)})}{n^{(n)} + 1}$  decreases with  $n$ !

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## Uncertainty about probability statements

smaller sets = more precise probability statements

### Lottery A

Number of winning tickets:  
exactly known as 5 out of 100

▶  $P(\text{win}) = 5/100$

### Lottery B

Number of winning tickets:  
not exactly known, supposedly  
between 1 and 7 out of 100

▶  $P(\text{win}) = [1/100, 7/100]$



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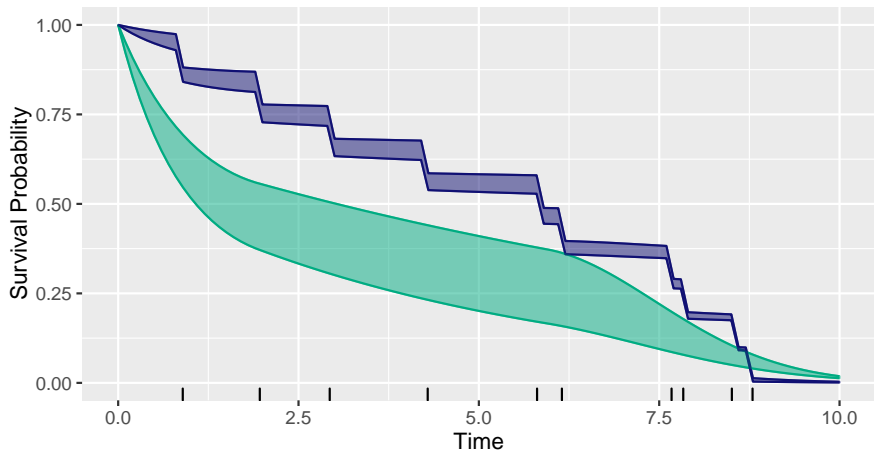
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  - ▶ Bounds for inferences (point estimate, prediction, ...)  
by min/max over

# Component Reliability with Sets of Priors

6/11



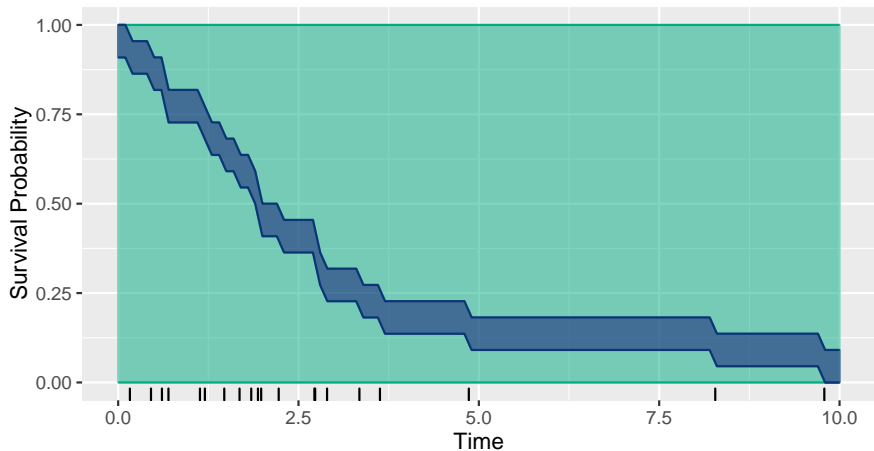
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6/11



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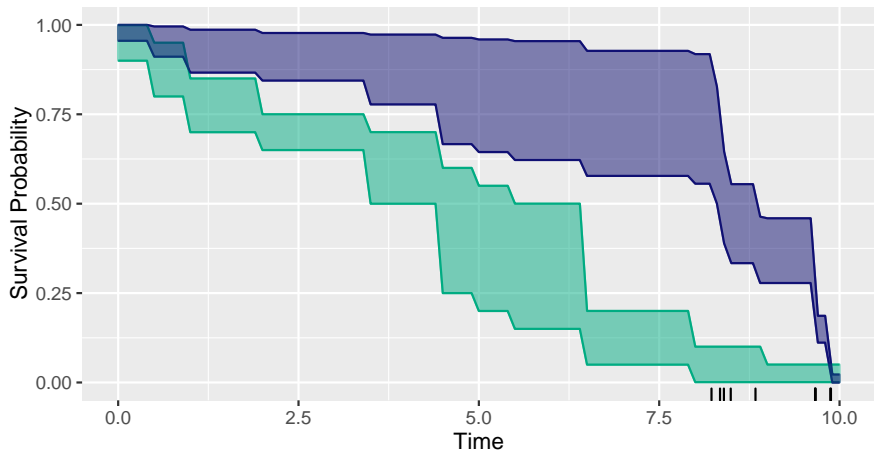
$$[\underline{y}^{(0)}, \bar{y}^{(0)}] = [0, 1]$$

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# Component Reliability with Sets of Priors

6/11



$$[\underline{n}^{(0)}, \bar{n}^{(0)}] = [1, 8]$$

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 Prior  Posterior

- ▶ Closed form for the system reliability via the survival signature:

$$\begin{aligned} R_{\text{sys}}(t \mid \bigcup_{k=1}^K \{n_{k,t}^{(0)}, y_{k,t}^{(0)}, t^k\}) &= P(T_{\text{sys}} > t \mid \dots) \\ &= \sum_{l_1=0}^{m_1} \dots \sum_{l_K=0}^{m_K} \Phi(l_1, \dots, l_K) \prod_{k=1}^K P(C_t^k = l_k \mid n_{k,t}^{(0)}, y_{k,t}^{(0)}, t^k) \end{aligned}$$

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Survival signature  $\Phi(l_1, \dots, l_K)$

(Coolen and Coolen-Maturi 2012)

$= P(\text{system functions} \mid \{l_k \text{ 's function}\}^{1:K})$

$l_1$	$l_2$	$l_3$	$\Phi$	$l_1$	$l_2$	$l_3$	$\Phi$
0	0	1	0	0	1	1	0
1	0	1	0	1	1	1	0
2	0	1	1/3	2	1	1	2/3
3	0	1	1	3	1	1	1
4	0	1	1	4	1	1	1

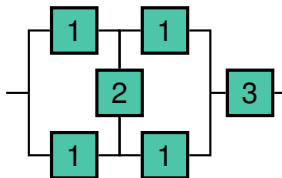
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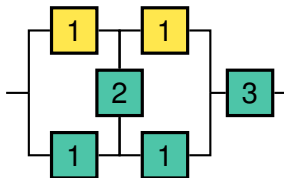
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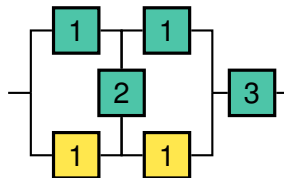
$$= \sum_{l_1=0}^{m_1} \dots \sum_{l_K=0}^{m_K} \Phi(l_1, \dots, l_K) \prod_{k=1}^K P(C_t^k = l_k \mid n_{k,t}^{(0)}, y_{k,t}^{(0)}, t^k)$$

Survival signature  $\Phi(l_1, \dots, l_K)$

(Coolen and Coolen-Maturi 2012)

$= P(\text{system functions} \mid \{l_k \text{ 'k' s function}\}^{1:K})$

$l_1$	$l_2$	$l_3$	$\Phi$	$l_1$	$l_2$	$l_3$	$\Phi$
0	0	1	0	0	1	1	0
1	0	1	0	1	1	1	0
2	0	1	1/3	2	1	1	2/3
3	0	1	1	3	1	1	1
4	0	1	1	4	1	1	1



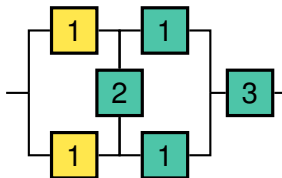
- Closed form for the system reliability via the survival signature:

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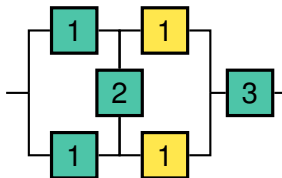
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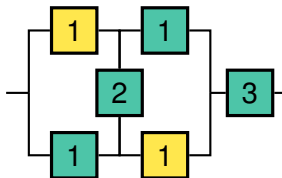
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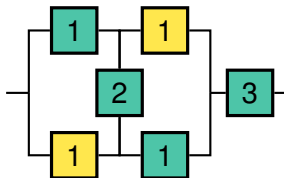
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Posterior predictive probability that in a new system,  $l_k$  of the  $m_k$  **k**'s function at time  $t$ :

$$\binom{m_k}{l_k} \int [P(T < t \mid p_t^k)]^{l_k} [P(T \geq t \mid p_t^k)]^{m_k - l_k} f(p_t^k \mid n_{k,t}^{(0)}, y_{k,t}^{(0)}, t^k) dp_t^k$$

- ▶ analytical solution for integral:  
 $C_t^k \mid n_{k,t}^{(0)}, y_{k,t}^{(0)}, t^k \sim \text{Beta-binomial}$

- ▶ Bounds for  $R_{\text{sys}}\left(t \mid \bigcup_{k=1}^K \{n_{k,t}^{(0)}, y_{k,t}^{(0)}, t^k\}\right)$  over  $\bigcup_{k=1}^K \{ \quad \}$ :
  - ▶  $\min R_{\text{sys}}(\cdot)$  by  $y_{k,t}^{(0)} = \underline{y}_{k,t}^{(0)}$  for any  $n_{k,t}^{(0)}$   
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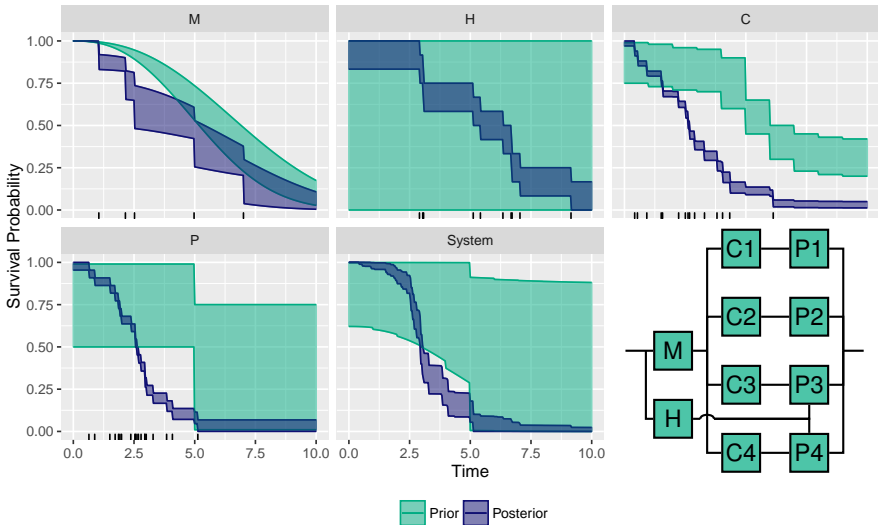
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  - ▶ implemented in **R** package `ReliabilityTheory` (Aslett 2016)

# System Reliability Bounds

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## Summary:

- ▶ Nonparametric modeling of component reliability curves
- ▶ Bayesian model combining expert knowledge and test data
- ▶ Set of system reliability functions reflects uncertainties from limited data, vague expert information, and prior-data conflict
- ▶ Easy-to-use implementation in **R** package  
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## Summary:

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## Next steps:

- ▶ Allow right-censored observations (RUL estimation)
- ▶ Allow dependence between components (common-cause failure, ...)
- ▶ Use for system design (where to put extra redundancy?)
- ▶ Use for maintenance planning

- Aslett, Louis (2016). *ReliabilityTheory: Tools for structural reliability analysis*. R package. URL: <http://www.louisaslett.com>.
- Coolen, Frank and Tahani Coolen-Maturi (2012). "Generalizing the Signature to Systems with Multiple Types of Components". In: *Complex Systems and Dependability*. Ed. by W. Zamojski et al. Vol. 170. Advances in Intelligent and Soft Computing. Springer, pp. 115–130. DOI: [10.1007/978-3-642-30662-4\\_8](https://doi.org/10.1007/978-3-642-30662-4_8).
- Evans, Michael and Hadas Moshonov (2006). "Checking for Prior-Data Conflict". In: *Bayesian Analysis* 1, pp. 893–914. URL: <http://projecteuclid.org/euclid.ba/1340370946>.
- Walter, Gero (2013). "Generalized Bayesian Inference under Prior-Data Conflict". PhD thesis. Department of Statistics, LMU Munich. URL: <http://edoc.ub.uni-muenchen.de/17059/>.
- Walter, Gero, Louis Aslett, and Frank Coolen (2016). "Bayesian Nonparametric System Reliability using Sets of Priors". Submitted to *International Journal of Approximate Reasoning*. URL: <http://arxiv.org/abs/1602.01650>.
- Walter, Gero and Thomas Augustin (2009). "Imprecision and Prior-data Conflict in Generalized Bayesian Inference". In: *Journal of Statistical Theory and Practice* 3, pp. 255–271. DOI: [10.1080/15598608.2009.10411924](https://doi.org/10.1080/15598608.2009.10411924).